

A Central Limit Theorem of Stationary Processes and the Parameter Estimation of Linear Processes

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§ 1 Some limit theorems on stationary processes

Let $\{Z(n); n \in J\}$ be a s -vector-valued linear process generated as $Z(n) = \sum_{j=0}^{\infty} G(j) e(n-j)$, $n \in J$, where the $Z(n)$'s have s components and the $e(n)$'s are p -vectors such that $E(e(n)) = 0$ and $E(e(n)e(n)') = \delta(m, n)K$ (K is a nonsingular $p \times p$ matrix); $G(j)$'s are $s \times p$ matrices; the components of Z , e , G are all real. If $\sum \text{tr} G(j)K G(j)' < \infty$, the process $\{Z(n)\}$ is a second order stationary process and has a spectral density matrix $f(\omega) = \frac{1}{2\pi} R(\omega)K R(\omega)'$, $-\pi \leq \omega \leq \pi$ where $R(\omega) = \sum_{j=0}^{\infty} G(j)e^{i\omega j}$. Denote by $C_Z(s)$ and $I_Z(\omega)$ respectively the serial covariance and the periodogram matrices constructed from a partial realization $\{Z(1), \dots, Z(N)\}$; namely, $C_Z(s) = \frac{1}{N} \sum_{m=1}^{N-s} Z(m)Z(m+s)'$, for $0 \leq s \leq N-1$, and $I_Z(\omega) =$

$C_z(t-s)$ for $-N+1 \leq s < 0$; $I_z(\omega) = F_z(\omega) F_z(\omega)^*$ where

$$F_z(\omega) = \frac{1}{\sqrt{2\pi N}} \sum_{n=1}^N z(n) e^{in\omega}$$

Denote the (α, β) component of $G(j)$, C_z and I_z by $G_{\alpha\beta}(j)$, $C_{\alpha\beta}^z$ and $C_{\alpha\beta}^z$ respectively and denote the α -th component of $z(n)$ and $e(n)$ by $z_\alpha(n)$ and $e_\alpha(n)$. Assuming that the process $\{e(n)\}$ is fourth-order stationary, let $G_{\alpha_1, \dots, \alpha_4}^e(t_1, \dots, t_4)$ be the fourth cumulant of $e_{\alpha_1}(t_1), \dots, e_{\alpha_4}(t_4)$ and assume that

$$\sum_{t_2, \dots, t_4 = -\infty}^{\infty} |G_{\alpha_1, \dots, \alpha_4}^e(0, t_2 - t_1, \dots, t_4 - t_1)| < \infty;$$

then the process $\{e(n)\}$ has a fourth-order spectral density $\tilde{G}_{\alpha_1, \dots, \alpha_4}^e(\omega_1, \omega_2, \omega_3)$ such that

$$\tilde{G}_{\alpha_1, \dots, \alpha_4}^e(\omega_1, \omega_2, \omega_3) = \frac{1}{(2\pi)^3} \sum_{t_1, \dots, t_3 = -\infty}^{\infty} \exp\{-i(\omega_1 t_1 + \dots + \omega_3 t_3)\} \\ \times G_{\alpha_1, \dots, \alpha_4}^e(0, t_1, t_2, t_3).$$

Denote by G_{g_1, \dots, g_4}^z and $\tilde{G}_{g_1, \dots, g_4}^z$ respectively the fourth-order cumulant and spectral density of the process $z(n)$,

Lemma 1.1. If $\sum_{j=0}^{\infty} |G_{\alpha\beta}(j)|^2 < \infty$ for each α, β and

$$\sum_{j_1, j_3 = -\infty}^{\infty} |G_{\alpha_1, \dots, \alpha_4}^e(j_1, j_2, j_3)| < \infty, \text{ the process}$$

$\{z(n)\}$ has a fourth-order spectral density

$$\tilde{G}_{g_1, \dots, g_4}^z(\omega_1, \omega_2, \omega_3) \text{ such that}$$

$$(1.1) \quad \hat{Q}_{g_1, \dots, g_4}^z(w_1, w_2, w_3) \\ = \sum_{d_1, \dots, d_4=1}^r k_{g_1, d_1}(w_1 + w_2 + w_3) k_{g_2, d_2}(-w_1) k_{g_3, d_3}(-w_2) k_{g_4, d_4}(w_3) \\ \times \hat{Q}_{d_1, \dots, d_4}^e(w_1 + w_2 + w_3, w_2, w_3)$$

Lemma 1.2. Assume $\sum_{j_1, j_2, j_3=-\infty}^{\infty} |Q_{d_1, \dots, d_4}^z(j_1, j_2, j_3)| < \infty$.

For any square-integrable functions W_1 and W_2 defined on

$$[-\pi, \pi], \quad \lim_{N \rightarrow \infty} N \text{Cov} \left\{ \int_{-\pi}^{\pi} W_1(\omega) I_{d_1, d_2}^z(\omega) d\omega, \int_{-\pi}^{\pi} W_2(\omega) I_{d_3, d_4}^z(\omega) d\omega \right\} \\ (1.2) \quad = 2\pi \int_{-\pi}^{\pi} W_1(\omega) \overline{W_2(\omega)} f_{d_1, d_3}(\omega) \overline{f_{d_2, d_4}(\omega)} d\omega \\ + 2\pi \int_{-\pi}^{\pi} W_1(\omega) \overline{W_2(\omega)} f_{d_1, d_4}(\omega) \overline{f_{d_2, d_3}(\omega)} d\omega \\ + 2\pi \iint_{-\pi}^{\pi} W_1(\omega) \overline{W_2(-\omega_2)} \hat{Q}_{d_1, \dots, d_4}^z(\omega, \omega_2, -\omega_2) d\omega, d\omega_2.$$

Lemma 1.3. If $\sum |G_{\alpha\beta}(j)|^2 < \infty$ for each α, β and $\sum_j |Q_{d_1, \dots, d_4}^e(j_1, j_2, j_3)| < \infty$,

$$\lim_{N \rightarrow \infty} N \text{Cov} \{ C_{d_1, d_2}^z(m), C_{d_3, d_4}^z(n) \} \\ = 2\pi \int_{-\pi}^{\pi} \left\{ f_{d_1, d_3}(\omega) \overline{f_{d_2, d_4}(\omega)} e^{-i(n-m)\omega} + f_{d_1, d_4}(\omega) \overline{f_{d_2, d_3}(\omega)} \right. \\ \left. \times e^{i(m+m)\omega} \right\} d\omega \\ + \sum_{\beta_1, \dots, \beta_r=1}^p \iint_{-\pi}^{\pi} \exp\{i(m\omega_1 + n\omega_2)\} k_{\alpha_1, \beta_1}(\omega_1) k_{\alpha_2, \beta_2}(-\omega) k_{\alpha_3, \beta_3}(\omega_2) k_{\alpha_4, \beta_4}(-\omega_2) \\ \times \hat{Q}_{\beta_1, \dots, \beta_4}^e(\omega_1, -\omega_2, \omega_2) d\omega_1, d\omega_2.$$

Let $\{W_m(n), \mathcal{F}_m(n); n=0, 1, \dots, n(m)\}$, $m=1, 2, \dots$, be a zero-mean square-integrable martingale for each m where $\{\mathcal{F}_m(n); n=1, 2, \dots, n(m)\}$ is a sequence of increasing σ -fields and $n(m) \rightarrow \infty$ as $m \rightarrow \infty$. Let $U_m(1) = W_m(1)$ and $U_m(k) = W_m(k)$ and $U_m(k) = W_m(k) - W_m(k-1)$ ($W_m(0) = 0$).

Lemma 1.4. [essentially due to Brown (1972)]. Suppose

that (i) $\lim_{m \rightarrow \infty} \frac{1}{n(m)} \sum_1^{n(m)} E[U_m(k)^2 I\{|U_m(k)| \geq \varepsilon n(m)\}] = 0$
for any $\varepsilon > 0$,

(ii) $\left[\sum_{k=1}^{n(m)} E\{U_m(k)^2 | \mathcal{F}_m(k-1)\} \right] / \sum_{k=1}^{n(m)} E\{U_m(k)^2\}$

tends to 1 in probability as $m \rightarrow \infty$; then

$\sum_1^{n(m)} U_m(k) / \sqrt{\sum E\{U_m(k)^2\}}$ is asymptotically normally distributed with mean 0 and variance 1.

Theorem 1.1. If a zero-mean vector-valued second order stationary process $\{x(t); t \in J\}$ satisfies that

(i) $\text{Var}\{E(x_\alpha(t+\tau) | \mathcal{F}_t)\} = O\left(\frac{1}{\tau^{2+\varepsilon}}\right)$ ($\varepsilon > 0$)

(ii) For a positive constant $\eta (> 0)$,

$$E\{E(x_\alpha(l)x_\beta(m) | \mathcal{F}_t) - E\{x_\alpha(l)x_\beta(m)\}\} \\ = O\left\{\frac{1}{(\min\{l-t, m-t\})^{1+\eta}}\right\} \quad l, m > t.$$

(iii) $\{x(t)\}$ has a spectral density matrix $f(\omega) = \{f_{\alpha\beta}(\omega)\}$ such that each element is continuous at the origin and $f(0)$ is non-degenerate, then $\frac{\sum_{n=1}^N x(n)}{\sqrt{N}}$ is asymptotically normally distributed with mean zero and covariance matrix $2\pi f(0)$, where \mathcal{F}_ε is the σ -field generated by the set of random vectors $\{x(n) : n \leq \varepsilon\}$.

Let $\{z(n)\}$ be the linear process $z(n) = \sum_{j=0}^{\infty} G(j)e(n-j)$, and denote by $\mathcal{B}(t)$ the σ -field generated by $\{e(m) : m \leq t\}$.

Theorem 1.2. Suppose (i) $\text{Var} \{ E \{ e_{\beta_1}(n) e_{\beta_2}(n+m) | \mathcal{B}(n-\tau) \} - \delta(m) k_{\beta_1, \beta_2} \} = O\left(\frac{1}{\tau^{2+\varepsilon}}\right)$ ($\varepsilon > 0$)

(ii) $E \{ E \{ e_{\beta_1}(n_1) e_{\beta_2}(n_2) e_{\beta_3}(n_3) e_{\beta_4}(n_4) | \mathcal{B}(n_1 - \tau) \} - E \{ e_{\beta_1}(n_1) e_{\beta_2}(n_2) e_{\beta_3}(n_3) e_{\beta_4}(n_4) \} \} = O\left(\frac{1}{\tau^{1+\eta}}\right)$

($\eta > 0, n_1 \leq n_2 \leq n_3 \leq n_4$),

(iii) $f_{\beta\beta}$ ($\beta = 1, \dots, s$) is square-integrable,

(iv) $\sum_{|j_1|, |j_2|, |j_3| = -\infty}^{\infty} |A_{\beta_1, \dots, \beta_4}^e(j_1, j_2, j_3)| < \infty$,

then $\sqrt{N} (C_{d_1 d_2}^z(m) - \delta_{d_1 d_2}^z(m))$ ($d_1, d_2 = 1, \dots, S$, $0 \leq m \leq L$) have a joint asymptotic normal distribution with mean 0 and the asymptotic covariance between $\sqrt{N} \{C_{d_1 d_2}^z(m_1) - \delta_{d_1 d_2}^z(m_1)\}$ and $\sqrt{N} \{C_{d_3 d_4}^z(m_2) - \delta_{d_3 d_4}^z(m_2)\}$ is given as

$$(1.4) \quad 2\pi \int_{-\pi}^{\pi} \{ f_{d_1 d_3}(\omega) \overline{f_{d_2 d_4}(\omega)} e^{-i(m_2 - m_1)\omega} + f_{d_1 d_4}(\omega) \overline{f_{d_2 d_3}(\omega)} e^{i(m_1 + m_2)\omega} \} d\omega$$

$$+ 2\pi \sum_{\beta_1, \dots, \beta_4=1}^p \iint_{-\pi}^{\pi} \exp \{ i m_1 \omega_1 + i m_2 \omega_2 \} K_{d_1 \beta_1}(\omega_1) K_{d_2 \beta_2}(-\omega_1)$$

$$\times K_{d_3 \beta_3}(\omega_2) K_{d_4 \beta_4}(-\omega_2) \hat{Q}_{\beta_1, \dots, \beta_4}^0(\omega_1, -\omega_2, \omega_2) d\omega_1 d\omega_2.$$

§2. Asymptotic properties of quasi-Gaussian maximum likelihood estimates for a linear process

Let P_0 be a family of spectral densities $f_0(\omega)$ $\theta \in \Theta \subset \mathbb{R}^b$. Then a functional T defined on P_0 is determined by the requirement that, for T , the set of all density such that $\sum_{i,j} (T_j)' < \infty$,

$$D(f_{T(f)}, f) = \min_{t \in \mathcal{E}} D(f_t, f)$$

where $D(f_t, f) = \int_{-\pi}^{\pi} \{ \log \det f_t(\omega) + \operatorname{tr} f_t(\omega)^{-1} f(\omega) \} d\omega$,

if there is a unique $T(f)$ in \mathcal{E} , let $f_N, f \in \mathcal{P}$ be such that for every continuous $s \times s$ matrix-valued function ψ

$$\int_{-\pi}^{\pi} \operatorname{tr} \psi(\omega) f_N(\omega) d\omega \rightarrow \int_{-\pi}^{\pi} \operatorname{tr} \psi(\omega) f(\omega) d\omega$$

as $N \rightarrow \infty$; then f_N is called weakly converge to f ($f_N \xrightarrow{w} f$).

Lemma 2.1. Suppose that \mathcal{E} is a compact subset of \mathbb{R}^s , $\theta_1 \neq \theta_2$ implies that $f_{\theta_1} \neq f_{\theta_2}$ on a set of positive definite, and that $f_{\theta}(\omega)$ is continuous in θ and ω . Then

(1) For every $f \in \mathcal{P}$, there is a value $T(f)$.

(2) Assume $T(f)$ is unique; then if $f_N \xrightarrow{w} f$ then

$$T(f_N) \rightarrow T(f).$$

(3) $T(f_{\theta}) = \theta$ uniquely for every $\theta \in \mathcal{E}$.

Lemma 2.2. Assume that every component of

$f_{\theta}(\omega)$ is twice continuously differentiable function of $\theta \in \Theta$ and that the second derivatives of these component are continuous in ω ; moreover, suppose that $T(f)$ exists uniquely and in $\text{Int}(\Theta)$ and that $M_f = \int_{-\pi}^{\pi} \left\{ \frac{\partial^2}{\partial \theta \partial \theta} \left(\text{tr } f_{\theta}(\omega)^{-1} f(\omega) \right) + \frac{\partial^2}{\partial \theta \partial \theta} \log \det f_{\theta}(\omega) \right\} d\omega$ $\theta = T(f)$ is a non-singular matrix. Then for every sequence of spectral density matrices $\{f_N\}$ such that $f_N \xrightarrow{w} f$;

$$(2.1) \quad T(f_N) = T(f) - \int_{-\pi}^{\pi} M_f^{-1} \frac{\partial}{\partial \theta} \left\{ \text{tr } f_{\theta}(\omega)^{-1} (f_N(\omega) - f(\omega)) \right\} d\omega \Big|_{\theta = T(f)} \\ + a_N \int_{-\pi}^{\pi} \frac{\partial}{\partial \theta} \left\{ \text{tr } f_{\theta}(\omega)^{-1} (f_N(\omega) - f(\omega)) \right\} d\omega \Big|_{\theta = T(f)}$$

where a_N is a $q \times q$ matrix which tends to zero as $N \rightarrow \infty$.

Lemma 2.3. Assume that $\{z(n)\}$ satisfies the conditions

(i) - (iv) of Theorem 1.2 and (v) $\sum_{j=-\infty}^{\infty} |j|^{\frac{1}{2}} |\gamma_{z^z}(j)| < \infty$.

Let $\phi_j(\omega)$, $j=1, \dots, q$ be $s \times s$ matrix-valued continuous functions on $[-\pi, \pi]$ such that $\phi_j(\omega) = \phi_j(\omega)^*$. Then,

$$(1) \quad p\text{-}\lim \int_{-\pi}^{\pi} \text{tr } I_z(\omega) \phi_j(\omega) d\omega = \int_{-\pi}^{\pi} \text{tr } f(\omega) \phi_j(\omega) d\omega.$$

(2) $\sqrt{N} \int_{-\pi}^{\pi} \text{tr} \{ I_z(\omega) - f(\omega) \} \phi_j(\omega) d\omega$, $j=1, \dots, g$ have asymptotically a normal distribution with zero mean-vector and covariance matrix V whose (j, l) element

is

$$(2.2) \quad 4\pi \int_{-\pi}^{\pi} \text{tr} f(\omega) \phi_j(\omega) f(\omega) \phi_l(\omega) d\omega + 2\pi \sum_{r,t,u,v=1}^g \iint_{-\pi}^{\pi} \phi_{rt}^{(j)}(\omega_1) \phi_{uv}^{(l)}(\omega_2) \tilde{R}_{rtuv}(-\omega_1, \omega_2, -\omega_2) d\omega_1 d\omega_2$$

where $\phi_{rt}^{(j)}(\omega)$ is the (r, t) th element of $\phi_j(\omega)$.

Theorem 2.1. Suppose that in addition to the conditions of Lemma 2.2, the conditions (i)-(iv) in Theorem 1.2 and (v) in Lemma 2.3 are satisfied.

Then p - $\lim_{N \rightarrow \infty} T(I_z) = T(f)$ and the vector $\sqrt{N}(T(I_z) - T(f))$,

under f , tends to the normal distribution

$N(0_g, M_f^{-1} \tilde{V} M_f^{-1})$ where $\tilde{V} = \{\tilde{V}_{j,l}\}$ is a $g \times g$ matrix such that

$$(2.3) \quad \tilde{V}_{j,l} = 4\pi \int_{-\pi}^{\pi} \text{tr} \left\{ f(\omega) \frac{\partial}{\partial \theta_j} f_0(\omega)^{-1} f(\omega) \frac{\partial}{\partial \theta_l} f_0(\omega)^{-1} \right\} d\omega$$

$$+ 2\pi \sum_{r,t,u,v=1}^g \iint_{-\pi}^{\pi} \frac{\partial}{\partial \theta_j} f_0^{(r,t)}(\omega_1)^{-1} \frac{\partial}{\partial \theta_l} f_0^{(u,v)}(\omega_2)^{-1} \tilde{R}_{rtuv}(\omega_1, \omega_2, -\omega_2) d\omega_1 d\omega_2$$

$-u_1, u_2, -u_2) du_1, du_2$

and $f_e^{(z, \tau)}(u)^{-1}$ is the (z, τ) th element of $f_e(u)^{-1}$.

Corollary

$$\tilde{V}_{je} = \pi \int_{-\pi}^{\pi} \text{tr} \left\{ f(u) \frac{\partial}{\partial \theta_j} f_e(u)^{-1} f(u) \frac{\partial}{\partial \theta_j} f_e(u)^{-1} \right\} \Big|_{\theta = \Pi(f)} du$$

$$+ \pi \sum_{a, b, c, d=1}^p \sum_{r, t, u, v=1}^s \int_{-\pi}^{\pi} \int_{-\pi}^{\pi} \frac{\partial}{\partial \theta_j} f_e^{(r, t)}(u_1)^{-1} \frac{\partial}{\partial \theta_j} f_e^{(u, v)}(u_2)^{-1} \Big|_{\theta = \Pi(f)}$$

$$\times K_{ra}(-u_1) K_{tb}(u_1) K_{uc}(-u_2) K_{vd}(u_2) \tilde{g}_{abcd}^e(-u_1, u_2, -u_2) du_1, du_2.$$

References

Alkhiezer, N.I. (1956), Theory of Approximation,
Frederick Ungar Publishing Co.

Brown, B.M. (1971), Martingale central limit theorems,
Ann. Math. Statist. 42, 59-66.

Dunsmuir, W. and Hannan, E.J. (1976), Vector linear
time series models, Adv. Appl. Prob. 8, 337-~~348~~
367.

Dunsmuir, W. (1979), A central limit theorem for
parameter estimation in stationary vector time
series and its application to models for a
signal observed whitewise, Ann. Stat., 7,

490-506.

- Fisher, R. A. (1925), Theory of statistical estimation, Proc. Cambridge Philos. Soc., 22, 702-725.
- Hannan, E. J. (1970). Multiple Time Series.
Wiley New York.
- Hannan, E. J. and Robinson, P. M. (1973). Lagged regression with unknown lags. J. R. S. S. 35, 252-267.
- Hannan, E. J. (1976). The asymptotic distribution of serial covariances. Ann. Statist. 4 376-399.
- Hosoya, Y. (1974). Estimation problems on stationary time-series models, Ph. D. thesis, Yale Univ.
- Ibragimov, I. A. and Linnik, Yu. (1971). Independent Stationary Sequences of Random Variables,
Walter Noordhoff Publishing.
- Rozanov, Yu. A. (1967), Stationary Random Processes,
Holden-Day, San Francisco.
- Taniguchi, M. (1979). On estimation of parameters of Gaussian stationary processes. J. Appl. Prob 16, 575-591.
- Walker, A. M. (1964). Asymptotic properties of least-square estimates of parameters of the spectrum

of a stationary nondeterministic time series,
J. Austral. Math. Soc. 4. 363-389.

Whittle, P. (1952), Some results in time series
analysis, Skandinavisk Aktuarietidskrift,
Hefte 1-2, 48-60.

Whittle, P. (1962), Gaussian estimation in
stationary time series, Bull. Inst. Int.
de Stat. tome 39-2, 105-129.