# Fluctuation of Population Size and <br> Effective Size of Population 

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## 集団の確率変動と有効個体数 <br> 飯塚 勝 <br> 九州歯科大学 数学教室


#### Abstract

The effective size of population has played an important role in population genetics．We consider a Wright－Fisher model whose population size is a simple Markov chain．For this model，we define inbreeding effective size and variance effective size．These effective sizes are turned out to be the same for this model．Effects of fluctuation of population size on the effective size are investigated．The effective size is not the same as the harmonic mean of population size unless fluctuation of population size is uncorrelated．The effective size is larger than the harmonic mean when the fluctuation of population size is positively autocorrelated and smaller than the harmonic mean when the fluctuation is negatively autocorrelated．


## 1．Effective size of population

In population genetics theory，a traditional formulation of simple stochastic haploid model is the following Wright－Fisher model．Consider a population of $N$ haploid individuals of which $i$ are of type $A_{1}$ and $N-i$ are of type $A_{2}$ ．The population reproduces itself in discrete generations．The whole of the next generation is formed by $N$ independent repetitions of the sampling with replacement．The probability that the next generation will contain $j$ members of type $A_{1}$ and $N-j$ of type $A_{2}$ is

$$
\begin{equation*}
p_{i j}=\binom{N}{j}\left(\frac{i}{N}\right)^{j}\left(1-\frac{i}{N}\right)^{N-j} . \tag{1}
\end{equation*}
$$

Let $Z(t)$ be the number of type $A_{1}$ in generation $t$ ．The process $\{Z(t)\}_{t=0,1,2, \ldots}$ is a dis－ crete time Markov chain on $\{0,1, \ldots, N-1, N\}$ with $p_{i j}=P(Z(t+1)=j \mid Z(t)=i)$ ．Let $x(t)=\frac{Z(t)}{N}$ be the gene frequency of type $A_{1}$ in generation $t$ in the population．The process $\{x(t)\}_{t=0,1,2, \ldots}$ is a discrete time Markov chain on $\left\{0, \frac{1}{N}, \ldots, \frac{N-1}{N}, 1\right\}$ with the transition matrix $T=\left(p_{i j}\right)_{i, j=0,1,2 \ldots, \ldots}$ ．This Markov chain is referred to as a Wright－Fisher model which has the following properties（Ethier and Kurtz（1986）and Ewens（1979））．The maximum eigenvalue of the transition matrix $T$ is 1 and the maximum non－unit eigenvalue is $\lambda_{2}=1-\frac{1}{N}$ ．The
probability that randomly sampled two genes from the population have the same parent is $\pi_{2}=\frac{1}{N}$. The quantity $\pi_{2}$ is referred to as the inbreeding coefficient. The conditional variance of $x(t+1)$ conditional on $x(t)=x$ is $\operatorname{Var}[x(t+1) \mid x(t)=x]=E\left[\{x(t+1)-x(t)\}^{2} \mid x(t)=x\right]$ $=\frac{x(1-x)}{N}$. The population size $N$ can be expressed as

$$
\begin{gather*}
N=\frac{1}{1-\lambda_{2}}  \tag{2}\\
N=\frac{1}{\pi_{2}} \tag{3}
\end{gather*}
$$

and

$$
\begin{equation*}
N=\frac{x(1-x)}{\operatorname{Var}[x(t+1) \mid x(t)=x]} \tag{4}
\end{equation*}
$$

For stochastic models that are more complicated than the Wright-Fisher model, the effective size of population plays an important role for the population size $N$ (Crow (1954), Kimura and Crow (1963) and Wright (1938)). There are several ways to define the effective size of population. The eigenvalue effective size $N e^{(e)}$ is defined by

$$
\begin{equation*}
N e^{(e)}=\frac{1}{1-\lambda_{2}} \tag{5}
\end{equation*}
$$

where $\lambda_{2}$ is the maximum non-unit eigenvalue of the transition matrix of the stochastic model. The inbreeding effective size $N e^{(i)}$ is defined by

$$
\begin{equation*}
N e^{(i)}=\frac{1}{\pi_{2}} \tag{6}
\end{equation*}
$$

where $\pi_{2}$ is the inbreeding coefficient of the stochastic model. The variance effective size $N e^{(v)}$ is defined by

$$
\begin{equation*}
N e^{(v)}=\frac{x(1-x)}{\operatorname{Var}[x(t+1) \mid x(t)=x]} \tag{7}
\end{equation*}
$$

where $x(t)$ is the gene frequency of type $A_{1}$ in generation $t$ of the stochastic model. For concrete examples of these effective sizes, see Crow and Kimura (1970), Ewens (1979) and Nagylaki (1992).

## 2. Effective size of fluctuating population

There is a lot of ecological data to the effect that the numbers of individuals in natural populations fluctuate considerably in each epoch and from generation to generation (see Andrewartha and Birch (1954), Elton and Nicholson (1942) and Odum (1959)). The variations in population size are influenced by such factors as climate, the abundance of available resources, fluctuation in prey-predator balance, competition with other species using the same habitat (Nicholson (1957)).

One of recent interests in theoretical population genetics is to find a mechanism for evaluating overdispersed molecular evolution (Gillespie (1991, 1993, 1994a, 1994b), Iwasa (1993), Ohta and Kimura (1971), Ohta and Tachida (1990), Tachida (1991) and Takahata (1987)). Overdispersed molecular evolution is a phenomena that the ratio of the variance
to the mean number of substitution of mutants among species is larger than one. This ratio is referred to as the dispersion index. Gillespie (1989), using data of 20 protein loci from three species of mammals, obtained 6.95 as an estimate of the dispersion index. Ohta (1995) analyzed 49 mammalian protein data and obtained an estimate of 5.6. Note that the dispersion index is equal to one for the neutral model since the substitution process is a Poisson process (Kimura (1983)). One of the candidates of the mechanism for evaluating the dispersion index being much larger than one is nearly neutral mutation model with fluctuating population size (Araki and Tachida (1997)).

Fluctuation of population size is not independent from generation to generation in general as in the case of stochastic selection (Guess and Gillespie (1977); Iizuka (1987), Iizuka and Matsuda (1982), Seno and Shiga (1984) and Takahata, Ishii and Matsuda (1975)). An appropriate concept for autocorrelated stochastic processes is mixing processes (Billingsley (1968)). In this paper, we will consider a simple case of mixing processes, that is, two-valued Markov chain as a model of autocorrelated fluctuation of population size. Our interest is how the effective size of population depends on the degree of autocorrelation of the fluctuation.

## 3. Model

Let $N(t)=N\left(t, \omega_{1}\right)$ be the population size in generation $t$. In this paper, we assume that $\left\{N\left(t, \omega_{1}\right)\right\}_{t=0, \pm 1, \pm 2, \ldots}$ is a stationary Markov chain on $\left\{N_{1}, N_{2}\right\}$ such that

$$
\begin{gather*}
P_{\omega_{1}}\left(N\left(t+1, \omega_{1}\right) \neq N\left(t, \omega_{1}\right) \mid N\left(t, \omega_{1}\right)=N_{j}\right)=q_{j},  \tag{8}\\
P_{\omega_{1}}\left(N\left(t+1, \omega_{1}\right)=N\left(t, \omega_{1}\right) \mid N\left(t, \omega_{1}\right)=N_{j}\right)=1-q_{j} \tag{9}
\end{gather*}
$$

and

$$
\begin{equation*}
P_{\omega_{1}}\left(N\left(0, \omega_{1}\right)=N_{j}\right)=p_{j}^{(s t)} \tag{10}
\end{equation*}
$$

where $1<N_{1}<N_{2}<\infty, 0 \leq q_{j} \leq 1, q_{1}+q_{2}>0$ and

$$
\begin{equation*}
\left(p_{1}^{(s t)}, p_{2}^{(s t)}\right)=\left(\frac{q_{2}}{q_{1}+q_{2}}, \frac{q_{1}}{q_{1}+q_{2}}\right) \tag{11}
\end{equation*}
$$

is the stationary distribution of the Markov chain. The parameter $q_{j}$ is the probability of changing size for $N_{j}(j=1,2)$.

Note that

$$
\begin{equation*}
\tau_{j}=\sum_{k=1}^{\infty} k\left(1-q_{j}\right)^{k-1} q_{j}=\frac{1}{q_{j}} \tag{12}
\end{equation*}
$$

is the mean persistence time for population size $N_{j}$ and

$$
\begin{equation*}
\tau=p_{1}^{(s t)} \tau_{1}+p_{2}^{(s t)} \tau_{2}=\frac{q_{1}^{2}+q_{2}^{2}}{q_{1} q_{2}\left(q_{1}+q_{2}\right)} \tag{13}
\end{equation*}
$$

is the mean persistence time of the Markov chain. The mean of $N\left(t, \omega_{1}\right)$ is

$$
\begin{equation*}
E_{\omega_{1}}\left[N\left(t, \omega_{1}\right)\right]=p_{1}^{(s t)} N_{1}+p_{2}^{(s t)} N_{2}=\frac{q_{2} N_{1}+q_{1} \dot{N}_{2}}{q_{1}+q_{2}} \tag{14}
\end{equation*}
$$

The covariance of $N\left(t, \omega_{1}\right)$ and $N\left(t+k, \omega_{1}\right)$ is

$$
\begin{equation*}
\operatorname{Cov}_{\omega_{1}}\left[N\left(t, \omega_{1}\right), N\left(t+k, \omega_{1}\right)\right]=\frac{q_{1} q_{2}\left(1-q_{1}-q_{2}\right)^{k}}{\left(q_{1}+q_{2}\right)^{2}}\left(N_{2}-N_{1}\right)^{2} . \tag{15}
\end{equation*}
$$

Eq.(15) means that constant population size, periodic change per generation, positively autocorrelated fluctuation, uncorrelated fluctuation, and negatively autocorrelated fluctuation correspond to $q_{1} q_{2}=0, q_{1}=q_{2}=1,0<q_{1}+q_{2}<1$ with $q_{1} q_{2} \neq 0, q_{1}+q_{2}=1$ with $q_{1} q_{2} \neq 0$, and $1<q_{1}+q_{2}<2$, respectively. The case of $q_{2}$ is much smaller than $q_{1}$ and $N_{1}$ is much smaller than $N_{2}$ corresponds to a model of population bottleneck since the mean persistence time $\tau_{1}$ for $N_{1}$ is much shorter than $\tau_{2}$ for $N_{2}$. The mean of $\frac{1}{N\left(t, \omega_{1}\right)}$ is

$$
\begin{equation*}
E_{\omega_{1}}\left[\frac{1}{N\left(t, \omega_{1}\right)}\right]=\frac{1}{N_{H}} \tag{16}
\end{equation*}
$$

where

$$
\begin{equation*}
N_{H}=\left\{\frac{p_{1}^{(s t)}}{N_{1}}+\frac{p_{2}^{(s t)}}{N_{2}}\right\}^{-1}=\left(q_{1}+q_{2}\right)\left(\frac{q_{2}}{N_{1}}+\frac{q_{1}}{N_{2}}\right)^{-1} \tag{17}
\end{equation*}
$$

is the harmonic mean of $N\left(t, \omega_{1}\right)$.
For fixed $\omega_{1}$, we consider a haploid population with type $A_{1}$ and $A_{2}$. The population size in generation $t$ is $N\left(t, \omega_{1}\right)$. The number of type $A_{1}$ in generation $t$ which is denoted by $Z\left(t, \omega_{1}, \omega_{2}\right)$ is a discrete time Markov process with

$$
\begin{align*}
& P_{\omega_{2}}\left(Z\left(t+1, \omega_{1}, \omega_{2}\right)=j \mid Z\left(t, \omega_{1}, \omega_{2}\right)=i\right) \\
& \quad=\binom{N\left(t+1, \omega_{1}\right)}{j}\left(\frac{i}{N\left(t, \omega_{1}\right)}\right)^{j}\left(1-\frac{i}{N\left(t, \omega_{1}\right)}\right)^{N\left(t+1, \omega_{1}\right)-j}, \tag{18}
\end{align*}
$$

$0 \leq i \leq N\left(t, \omega_{1}\right), 0 \leq j \leq N\left(t+1, \omega_{1}\right)$. Let $x(t)=x\left(t, \omega_{1}, \omega_{2}\right)=\frac{Z\left(t, \omega_{1}, \omega_{2}\right)}{N\left(t, \omega_{1}\right)}$ be the gene frequency of type $A_{1}$ in generation $t$. The process $\left\{x\left(t, \omega_{1}, \omega_{2}\right)\right\}_{t=0,1,2 \ldots} \ldots$ is a Wright-Fisher model with variable population size $\left\{N\left(t, \omega_{1}\right)\right\}$ for fixed $\omega_{1}$. Incorporating stochastic effects by fluctuating population size, this model is referred to as a Wright-Fisher model with fluctuating population size. Note that $x\left(t, \omega_{1}, \omega_{2}\right)$ is $\sigma\left(x\left(t-1, \omega_{1}, \omega_{2}\right), N\left(t-1, \omega_{1}\right), N\left(t, \omega_{1}\right)\right)$ measurable and $N\left(t, \omega_{1}\right)$ is $\sigma\left(N\left(s, \omega_{1}\right), s \leq t-1\right)$ measurable where $\sigma\left(x\left(t-1, \omega_{1}, \omega_{2}\right), N\left(t-1, \omega_{1}\right), N\left(t, \omega_{1}\right)\right)$ is a $\sigma$-field generated by $x\left(t-1, \omega_{1}, \omega_{2}\right), N\left(t-1, \omega_{1}\right), N\left(t, \omega_{1}\right)$ and $\sigma\left(N\left(s, \omega_{1}\right), s \leq t-1\right)$ is a $\sigma$-field generated by $N\left(s, \omega_{1}\right), s \leq t-1$.

Because of the autocorrelation of fluctuation of population size, we must extend the definition of effective size, which will be done for inbreeding effective size and variance effective size in the following. Let $\pi_{2}(t)$ be the probability that randomly sampled two genes from the population in generation $t$ have the same ancestral gene. We have

$$
\begin{equation*}
1-\pi_{2}(t)=E_{\omega_{1}}\left[\prod_{k=0}^{t-1}\left(1-\frac{1}{N\left(k, \omega_{1}\right)}\right)\right]\left\{1-\pi_{2}(0)\right\} \tag{19}
\end{equation*}
$$

Since

$$
\begin{equation*}
1-\pi_{2}(t)=\left(1-\frac{1}{N}\right)^{t}\left\{1-\pi_{2}(0)\right\} \tag{20}
\end{equation*}
$$

in the case of constant population size $\left(N\left(t, \omega_{1}\right)=N\right)$, the inbreeding effective size $N e^{(i)}=$ $N e^{(i)}\left(q_{1}, q_{2}\right)$ can be defined by

$$
\begin{equation*}
\left.1-\frac{1}{N e^{(i)}}=\lim _{t \rightarrow \infty}\left\{E_{\omega_{1}} \prod_{k=0}^{t-1}\left(1-\frac{1}{N\left(k, \omega_{1}\right)}\right)\right]\right\}^{\frac{1}{t}} \tag{21}
\end{equation*}
$$

Let $V_{t}(x)$ be the conditional variance of $x\left(t, \omega_{1}, \omega_{2}\right)$ conditional on $x\left(0, \omega_{1}, \omega_{2}\right)=x$, that is,

$$
\begin{align*}
V_{t}(x) & =\operatorname{Var}_{\left(\omega_{1}, \omega_{2}\right)}\left[x\left(t, \omega_{1}, \omega_{2}\right) \mid x\left(0, \omega_{1}, \omega_{2}\right)=x\right] \\
& =E_{\left(\omega_{1}, \omega_{2}\right)}\left[\left\{x\left(t, \omega_{1}, \omega_{2}\right)-x\left(0, \omega_{1}, \omega_{2}\right)\right\}^{2} \mid x\left(0, \omega_{1}, \omega_{2}\right)=x\right] . \tag{22}
\end{align*}
$$

Since

$$
\begin{equation*}
V_{t}(x)=\left\{1-\left(1-\frac{1}{N}\right)^{t}\right\} x(1-x) \tag{23}
\end{equation*}
$$

in the case of constant population size $\left(N\left(t, \omega_{1}\right)=N\right)$, the variance effective size $N e^{(v)}=$ $N e^{(v)}\left(q_{1}, q_{2}\right)$ can be defined by

$$
\begin{equation*}
1-\frac{1}{N e^{(v)}}=\lim _{t \rightarrow \infty}\left\{1-\frac{V_{t}(x)}{x(1-x)}\right\}^{\frac{1}{t}} \tag{24}
\end{equation*}
$$

We can show that the inbreeding effective size is the same as the variance effective size for this model.

Lemma 1 For this model,

$$
\begin{equation*}
N e^{(i)}=N e^{(v)} \tag{25}
\end{equation*}
$$

Proof. It is enough to show that

$$
\begin{equation*}
\left.1-\frac{1}{N e^{(v)}}=\lim _{t \rightarrow \infty}\left\{E_{\omega_{1}} \prod_{k=0}^{t-1}\left(1-\frac{1}{N\left(k, \omega_{1}\right)}\right)\right]\right\}^{\frac{1}{t}} \tag{26}
\end{equation*}
$$

Taking a conditional expectation, we have

$$
\begin{align*}
& E_{\left(\omega_{1}, \omega_{2}\right)}\left[x\left(t, \omega_{1}, \omega_{2}\right)\left\{1-x\left(t, \omega_{1}, \omega_{2}\right)\right\} \mid x\left(0, \omega_{1}, \omega_{2}\right)=x\right] \\
& \quad=E_{\left(\omega_{1}, \omega_{2}\right)}\left[E_{\omega_{2}}\left[x\left(t, \omega_{1}, \omega_{2}\right)\left\{1-x\left(t, \omega_{1}, \omega_{2}\right)\right\} \mid N\left(t-1, \omega_{1}\right)\right] \mid x\left(0, \omega_{1}, \omega_{2}\right)=x\right] \\
& \quad=E_{\left(\omega_{1}, \omega_{2}\right)}\left[\left.\left(1-\frac{1}{N\left(t-1, \omega_{1}\right)}\right) x\left(t-1, \omega_{1}, \omega_{2}\right)\left\{1-x\left(t-1, \omega_{1}, \omega_{2}\right)\right\} \right\rvert\, x(0)=x\right] \tag{27}
\end{align*}
$$

Iterating this operation, we have

$$
\begin{equation*}
\left.E_{\left(\omega_{1}, \omega_{2}\right)}\left[x\left(t, \omega_{1}, \omega_{2}\right)\left\{1-x\left(t, \omega_{1}, \omega_{2}\right)\right\} \mid x\left(0, \omega_{1}, \omega_{2}\right)=x\right]=E_{\omega_{1}} \prod_{k=0}^{t-1}\left(1-\frac{1}{N\left(k, \omega_{1}\right)}\right)\right] x(1-x) \tag{28}
\end{equation*}
$$

Since $\left\{x\left(t, \omega_{1}, \omega_{2}\right)\right\}$ is a martingale $\left(E_{\left(\omega_{1}, \omega_{2}\right)}\left[x\left(t, \omega_{1}, \omega_{2}\right) \mid x\left(0, \omega_{1}, \omega_{2}\right)=x\right]=x\right)$,

$$
\begin{align*}
& E_{\left(\omega_{1}, \omega_{2}\right)}\left[\left\{x\left(t, \omega_{1}, \omega_{2}\right)-x\left(0, \omega_{1}, \omega_{2}\right)\right\}^{2} \mid x\left(0, \omega_{1}, \omega_{2}\right)=x\right] \\
& \quad=-E_{\left(\omega_{1}, \omega_{2}\right)}\left[x\left(t, \omega_{1}, \omega_{2}\right)\left\{1-x\left(t, \omega_{1}, \omega_{2}\right)\right\} \mid x\left(0, \omega_{1}, \omega_{2}\right)=x\right]+x(1-x) \\
& \quad=\left\{1-E_{\omega_{1}}\left[\prod_{k=0}^{t-1}\left(1-\frac{1}{N\left(k, \omega_{1}\right)}\right)\right]\right\} x(1-x) . \tag{29}
\end{align*}
$$

We have the conclusion by the definition of the variance effective size (Eq.(24)).
We will use a notation $N e=N e\left(q_{1}, q_{2}\right)$ for $N e^{(i)}$ and $N e^{(v)}$. Our interests are as follows.
Is the effective size $N e$ equal to the harmonic mean $N_{H}$ ?
How does the effective size $N e\left(q_{1}, q_{2}\right)$ depend on $q_{1}$ and $q_{2}$ ?
We will consider these problems in the next section. For models with fluctuating population, see Chia and Pollak (1974), Donnelly (1986), Heyde and Seneta (1975), Karlin (1968), Klebaner (1988) and Seneta (1974).

## 4. Results

First, we will present a concrete expression for the effective size $N e$. For this end, we prepare two lemmas.

Lemma 2 For $i=1,2$, the conditional expectation

$$
\begin{equation*}
B_{i}(t)=E_{\omega_{1}}\left[\left.\prod_{k=0}^{t-1}\left(1-\frac{1}{N\left(k, \omega_{1}\right)}\right) \right\rvert\, N\left(0, \omega_{1}\right)=N_{i}\right] \tag{30}
\end{equation*}
$$

satisfies

$$
\begin{align*}
& B_{i}(t+2)-\left\{\left(1-\frac{1}{N_{1}}\right)\left(1-q_{1}\right)+\left(1-\frac{1}{N_{2}}\right)\left(1-q_{2}\right)\right\} B_{i}(t+1) \\
& \quad+\left(1-\frac{1}{N_{1}}\right)\left(1-\frac{1}{N_{2}}\right)\left(1-q_{1}-q_{2}\right) B_{i}(t)=0 \tag{31}
\end{align*}
$$

Proof. For $i, j=1,2(i \neq j)$, we have

$$
\begin{align*}
& B_{i}(t+1)=E_{\omega_{1}}\left[\left.\prod_{k=0}^{t}\left(1-\frac{1}{N\left(k, \omega_{1}\right)}\right) \right\rvert\, N\left(1, \omega_{1}\right)=N_{i}\right] P\left(N\left(1, \omega_{1}\right)=N_{i} \mid N\left(0, \omega_{1}\right)=N_{i}\right) \\
& \quad+E_{\omega_{1}}\left[\left.\prod_{k=0}^{t}\left(1-\frac{1}{N\left(k, \omega_{1}\right)}\right) \right\rvert\, N\left(1, \omega_{1}\right)=N_{j}\right] P\left(N\left(1, \omega_{1}\right)=N_{j} \mid N\left(0, \omega_{1}\right)=N_{i}\right) \\
& \quad=\left(1-\frac{1}{N_{i}}\right)\left\{\left(1-q_{i}\right) B_{i}(t)+q_{i} B_{j}(t)\right\} \tag{32}
\end{align*}
$$

which implies the conclusion.
Let $\alpha_{+}$and $\alpha_{-}\left(\alpha_{+} \geq \alpha_{-}\right)$be solutions to

$$
\begin{align*}
f(\alpha) & =\alpha^{2}-\left\{\left(1-\frac{1}{N_{1}}\right)\left(1-q_{1}\right)+\left(1-\frac{1}{N_{2}}\right)\left(1-q_{2}\right)\right\} \alpha \\
& +\left(1-\frac{1}{N_{1}}\right)\left(1-\frac{1}{N_{2}}\right)\left(1-q_{1}-q_{2}\right)=0 . \tag{33}
\end{align*}
$$

Note that $\alpha_{+}$and $\alpha_{-}$are real and $\alpha_{+}=\alpha_{-}$if and only if $1-\frac{1}{N_{1}}=\left(1-\frac{1}{N_{2}}\right)\left(1-q_{2}\right)$.

Lemma 3 For $q_{1}=0$ and $1-\frac{1}{N_{1}}=\left(1-\frac{1}{N_{2}}\right)\left(1-q_{2}\right)$,

$$
\begin{equation*}
E_{\omega_{1}}\left[\prod_{k=0}^{t-1}\left(1-\frac{1}{N\left(k, \omega_{1}\right)}\right)\right]=\left(1-\frac{1}{N_{1}}\right)^{t} \tag{34}
\end{equation*}
$$

and

$$
\begin{equation*}
E_{\omega_{1}}\left[\prod_{k=0}^{t-1}\left(1-\frac{1}{N\left(k, \omega_{1}\right)}\right)\right]=\frac{c_{1} q_{2}+c_{2} q_{1}}{q_{1}+q_{2}} \alpha_{+}^{t-1}+\frac{d_{1} q_{2}+d_{2} q_{1}}{q_{1}+q_{2}} \alpha_{-}^{t-1} \tag{35}
\end{equation*}
$$

otherwise. Here

$$
\begin{equation*}
c_{i}=\frac{1-\frac{1}{N_{i}}}{\alpha_{+}-\alpha_{-}}\left\{\left(1-\frac{1}{N_{i}}\right)\left(1-q_{i}\right)+\left(1-\frac{1}{N_{j}}\right) q_{i}-\alpha_{-}\right\} \tag{36}
\end{equation*}
$$

and

$$
\begin{equation*}
d_{i}=-\frac{1-\frac{1}{N_{i}}}{\alpha_{+}-\alpha_{-}}\left\{\left(1-\frac{1}{N_{i}}\right)\left(1-q_{i}\right)+\left(1-\frac{1}{N_{j}}\right) q_{i}-\alpha_{+}\right\} \tag{37}
\end{equation*}
$$

$i, j=1,2$ and $i \neq j$.
Proof. Since

$$
\begin{equation*}
E_{\omega_{1}}\left[\prod_{k=0}^{t-1}\left(1-\frac{1}{N\left(k, \omega_{1}\right)}\right)\right]=B_{1}(t) p_{1}^{(s t)}+B_{2}(t) p_{2}^{(s t)} \tag{38}
\end{equation*}
$$

we have the conclusion by solving the recurrence equation for $B_{1}(t)$ and $B_{2}(t)$.
The following theorem presents a concrete expression for the effective size.

Theorem 1 For $q_{1} q_{2} \neq 0$,

$$
\begin{equation*}
N e=\frac{1}{1-\alpha_{+}} \tag{39}
\end{equation*}
$$

For $q_{j}=0$,

$$
\begin{equation*}
N e=N_{H}=N_{j} . \tag{40}
\end{equation*}
$$

Proof. Note that for real numbers $a, b, A$ and $B$ with $A>|B|, a>0$

$$
\begin{equation*}
\lim _{t \rightarrow \infty}\left(a A^{t}+b B^{t}\right)^{\frac{1}{t}}=A \tag{41}
\end{equation*}
$$

and for for real numbers $c, d$ and $C$ with $C>0, c>|d|$,

$$
\begin{equation*}
\lim _{t \rightarrow \infty}\left\{c C^{t}+d(-C)^{t}\right\}^{\frac{1}{t}}=C \tag{42}
\end{equation*}
$$

We have the conclusion by Lemma 3 .
Next, we consider the relation between the effective size $N e$ and the harmonic mean $N_{H}$. From the sign of $f\left(1-\frac{1}{N_{1}}\right), f\left(1-\frac{1}{N_{2}}\right)$ and $f\left(1-\frac{1}{N_{H}}\right)$, we can obtain the following result (Note that $f\left(1-\frac{1}{N e}\right)=0$ ).

Theorem 2 For positively autocorrelated fluctuation $\left(0<q_{1}+q_{2}<1\right.$ and $\left.q_{1} q_{2} \neq 0\right)$,

$$
\begin{equation*}
N_{1}<N_{H}<N e<N_{2} \tag{43}
\end{equation*}
$$

For uncorrelated fluctuation $\left(q_{1}+q_{2}=1\right.$ and $\left.q_{1} q_{2} \neq 0\right)$,

$$
\begin{equation*}
N_{1}<N_{H}=N e<N_{2} . \tag{44}
\end{equation*}
$$

For negatively autocorrelated fluctuation ( $q_{1}+q_{2}>1$ ),

$$
\begin{equation*}
N_{1}<N e<N_{H}<N_{2} . \tag{45}
\end{equation*}
$$

By this result, the effective size is equal to the harmonic mean if and only if the fluctuation of population size is uncorrelated.

We consider the dependence of $q_{1}$ and $q_{2}$ on the effective size $N e$. Differentiating $\alpha_{+}=\alpha_{+}\left(q_{1}, q_{2}\right)$ by $q_{1}$ and $q_{2}$, we can obtain the following result.

Theorem 3 For fixed $q_{2}\left(q_{2} \neq 0\right), N e=N e\left(q_{1}, q_{2}\right)$ is an increasing function of $q_{1}$. For fixed $q_{1}\left(q_{1} \neq 0\right), N e=N e\left(q_{1}, q_{2}\right)$ is a decreasing function of $q_{2}$.

In the next theorem, we consider the case where $q_{2}$ is proportional to $q_{1}$. We can obtain the following result in the same way as Theorem 3.

Theorem 4 For fixed $c(c>0)$, we set $q_{1}=q$ and $q_{2}=c q$. For $0<q<\min \left\{1, \frac{1}{c}\right\}$, $N e=N e(q)$ is a decreasing function of $q$ for fixed $c$.
J. H. Gillespie performed computer simulations for the case of $c=1$ where mutation and selection are incorporated. He found that average heterozygosity (a measure of genetic diversity) is an increasing function of the mean persistence time $\tau$. Since the values of parameters in his computer simulations are restricted, he is interested in whether this is a general phenomena or not (Gillespie:personal communication). Theorem 4 is consistent with this phenomena, since $\tau=\frac{1}{q}$ when $c=1$ and average heterozygosity is an increasing function of $N e$. This means that his finding by computer simulations seems to be a general phenomena. Indeed, this is a motivation of the present paper.

Theorem 4 implies that the weaker the autocorrelation of fluctuation of population size is, the smaller the effective size is. An explanation of this result is as follows. When the autocorrelation is weak, it is difficult to predict what will happen to changes in population size in the next several generation. It may be deleterious to the population. On the other hand, when the effective size is small, the population has little genetic variation which may cause the extinction of the population by a sudden environmental change that has a harmful effect to the majority (wild type) of individuals in the population.

## 5. Asymptotic relations

The sizes of population $N_{1}$ and $N_{2}$ may be very large in natural populations. Further, the autocorrelation of $\left\{N\left(t, \omega_{1}\right)\right\}$ may be very strong. In such cases, we can obtain asymptotic behavior of the effective size. For this end, we parameterize $N_{1}, N_{2}, N e, N_{H}, q_{1}$ and $q_{2}$ by $\varepsilon$ such as $N_{1}^{\varepsilon}, N_{2}^{\varepsilon}, N e^{\varepsilon}, N_{H}^{\varepsilon}, q_{1}^{\varepsilon}$ and $q_{2}^{\varepsilon}\left(N_{1}^{\varepsilon}<N_{2}^{\varepsilon}\right)$.

In the following theorem, we consider the case where $N_{1}^{\varepsilon}$ and $N_{2}^{\varepsilon}$ are very large and the ratio of $N_{2}^{\varepsilon}$ to $N_{1}^{\varepsilon}$ is finite.

Theorem 5 Assume that $\lim _{\varepsilon \rightarrow 0} N_{1}^{\varepsilon}=\infty$ with $\lim _{\varepsilon \rightarrow 0} \frac{N_{1}^{\varepsilon}}{N_{2}^{\varepsilon}}>0$, then

$$
\begin{equation*}
\lim _{\varepsilon \rightarrow 0} \frac{N e^{\varepsilon}}{N_{2}^{\varepsilon}}=1 \tag{46}
\end{equation*}
$$

if $\lim _{\varepsilon \rightarrow 0} N_{1}^{\varepsilon}\left(q_{1}^{\varepsilon}+q_{2}^{\varepsilon}\right)=0$ and

$$
\begin{equation*}
\lim _{\varepsilon \rightarrow 0} \frac{N e^{\varepsilon}}{N_{H}^{\varepsilon}}=1 \tag{47}
\end{equation*}
$$

if $\lim _{\varepsilon \rightarrow 0} N_{1}^{\varepsilon}\left(q_{1}^{\varepsilon}+q_{2}^{\varepsilon}\right)=\infty$.

The following result is more general than Theorem 5, since it is not necessarily assumed that $N_{1}^{\varepsilon}$ and $N_{2}^{\varepsilon}$ are very large.

Theorem 6 Assume that $\lim _{\varepsilon \rightarrow 0} N_{1}^{\varepsilon}\left(q_{1}^{\varepsilon}+q_{2}^{\varepsilon}\right)=0$ and $0<\lim _{\varepsilon \rightarrow 0} \frac{N_{1}^{\varepsilon}}{N_{2}^{\varepsilon}}<1$, then

$$
\begin{equation*}
\lim _{\varepsilon \rightarrow 0} \frac{N e^{\varepsilon}}{N_{2}^{\varepsilon}}=1 \tag{48}
\end{equation*}
$$

Assume that $\lim _{\varepsilon \rightarrow 0} N_{2}^{\varepsilon}\left(q_{1}^{\varepsilon}+q_{2}^{\varepsilon}\right)=0$ and $\lim _{\varepsilon \rightarrow 0} \frac{N_{1}^{\varepsilon}}{N_{2}^{\varepsilon}}=0$, then

$$
\begin{equation*}
\lim _{\varepsilon \rightarrow 0} \frac{N e^{\varepsilon}}{N_{2}^{\varepsilon}}=1 \tag{49}
\end{equation*}
$$

Assume that $\lim _{\varepsilon \rightarrow 0} N_{1}^{\varepsilon}\left(q_{1}^{\varepsilon}+q_{2}^{\varepsilon}\right)=0$ and $0<\lim _{\varepsilon \rightarrow 0} N_{2}^{\varepsilon}\left(q_{1}^{\varepsilon}+q_{2}^{\varepsilon}\right) \leq \infty$, then

$$
\begin{equation*}
\lim _{\varepsilon \rightarrow 0} \frac{N e^{\varepsilon}}{N_{2}^{\varepsilon}}=\frac{1}{1+\lim _{\varepsilon \rightarrow 0} N_{2}^{\varepsilon} q_{2}^{\varepsilon}} \tag{50}
\end{equation*}
$$

Assume that $\lim _{\varepsilon \rightarrow 0} N_{1}^{\varepsilon}\left(q_{1}^{\varepsilon}+q_{2}^{\varepsilon}\right)=\infty$, then

$$
\begin{equation*}
\lim _{\varepsilon \rightarrow 0} \frac{N e^{\varepsilon}}{N_{H}^{\varepsilon}}=1 \tag{51}
\end{equation*}
$$

The autocorrelation of fluctuation of population size can be classified as follows. The case of $\lim _{\varepsilon \rightarrow 0} N_{2}^{\varepsilon}\left(q_{1}^{\varepsilon}+q_{2}^{\varepsilon}\right)=0$ is referred to as strong autocorrelation. The case of $\lim _{\varepsilon \rightarrow 0} N_{1}^{\varepsilon}\left(q_{1}^{\varepsilon}+q_{2}^{\varepsilon}\right)=\infty$
is referred to as weak autocorrelation. The other cases are classified as moderate autocorrelation (See Gillespie (1991) for such a classification in stochastic selection models). Theorem 5 and Theorem 6 imply that the effective size is very close to the harmonic mean when the fluctuation has weak autocorrelation. On the other hand, it is very close to the larger population size when the fluctuation has strong autocorrelation.

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