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A strong law of large numbers for branching processes: almost sure spine events

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Abstract

We demonstrate a novel strong law of large numbers for branching processes, with a simple proof via measure-theoretic manipulations and spine theory. Roughly speaking, any sequence of events that eventually occurs almost surely for the spine entails the almost sure convergence of a certain sum over particles in the population.

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1 Introduction

We shall work with a fairly general Markov branching process. To define this process, we suppose that we are given three ingredients:

- A Markov process ψ_t , $t \ge 0$, in a measurable space (J, \mathcal{B}) ;
- A measurable function $R: J \to [0,\infty)$;
- A collection of random variables A(x), $x \in J$ taking values in $\{0, 1, 2, ...\}$, such that $M(x) := E[A(x)] 1 < \infty$.

Our branching process is then defined, under a probability measure \mathbb{P} , as follows: we begin with one particle. This particle moves around in J like a copy of the process ψ_t . When at position x, it dies at rate R(x), that is, if \emptyset is our original particle, $X_{\emptyset}(t)$ is its position at time t and τ_{\emptyset} is its time of death, then

$$\mathbb{P}(\tau_{\emptyset} > t \mid X_{\emptyset}(s), \ s \le t) = \exp\left(-\int_{0}^{t} R(X_{\emptyset}(s))ds\right)$$

At its time of death, it is replaced in its position x by a random number of children, the number being specified by a copy of A(x). These children then each independently show the same stochastic behaviour as their parent, moving around like independent copies of ψ_t and branching at rate R(x) when at position x into a random number of particles that is an independent copy of A(x). We let N(t) be the set of all particles that are alive at time t; if $v \in N(t)$ then we let $X_v(t)$ be the position of particle v at time t; and we let A_v be the number of children of particle v.

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We let \mathcal{F}_t , $t \geq 0$ be the natural filtration of this process. We now extend our probability measure \mathbb{P} to a new probability measure $\tilde{\mathbb{P}}$ on a bigger space by choosing one special line of descent which we call the *spine* as follows. The initial particle is part of the spine, and when a spine particle dies the new spine particle is chosen uniformly from amongst its children. We let the natural filtration of the new process, in which there is a branching process with one marked line of descent, be $\tilde{\mathcal{F}}_t$, $t \geq 0$. Let $\operatorname{spine}(t)$ be the spine particle at time t, and ξ_t be its position; let S(t) be the set of particles that have been in the spine up to time t, $S(t) = {\operatorname{spine}(s) : s \leq t}$.

For details of all of the above, see [1] or Chapter 2 of [3].

Note that if X is an \mathcal{F}_t -measurable random variable, then there exist \mathcal{F}_t -measurable random variables such that we may write

$$X = \sum_{v \in N(t)} X_v \mathbb{1}_{\{\operatorname{spine}(t)=v\}}.$$

The reason for this representation is essentially that if X is $\tilde{\mathcal{F}}_t$ -measurable, then $X = X(\text{spine}(t), \{X_u(s), u \in N(t), s \leq t\})$ and we can choose $X_v = X(v, \{X_u(s), u \in N(t), s \leq t\})$. This can be made rigorous: see page 24 of [3].

Suppose that $\zeta(t)$ is a non-negative martingale with respect to the filtration $\mathcal{G}_t := \sigma(\xi_s, s \leq t)$, such that $\tilde{\mathbb{E}}[\zeta(t)] = 1$. We may write

$$\zeta(t) = \sum_{v \in N(t)} \zeta_v(t) \mathbb{1}_{\{\text{spine}(t)=v\}}$$

where each $\zeta_v(t)$ is an \mathcal{F}_t -measurable random variable. Then

$$\tilde{\zeta}(t) := e^{-\int_0^t M(\xi_s) R(\xi_s) ds} \zeta(t) \prod_{v \in S(t)} (1 + A_v)$$

is a martingale with respect to $\tilde{\mathcal{F}}_t$ (see Theorem 2.4 of [3]). We define a new measure $\tilde{\mathbb{Q}}$ by setting

$$\left. \frac{d\tilde{\mathbb{Q}}}{d\tilde{\mathbb{P}}} \right|_{\tilde{\mathcal{F}}_t} := \tilde{\zeta}(t)$$

The measure $\hat{\mathbb{Q}}$ has a nice description in terms of the spine, although this will not be used in this article. Briefly, the motion of the spine is biased by the martingale $\zeta(t)$; branching events along the spine occur at an accelerated rate $(1 + M(\xi_t))R(\xi_t)$ when the spine is at position ξ_t ; and the number of children of the spine is size-biased. All other (non-spine) particles, once born, remain unaffected.

We also let $\mathbb{Q} := \tilde{\mathbb{Q}}|_{\mathcal{F}_t}$ be a measure on \mathcal{F}_t , the natural filtration of the original branching process. Then

$$\left. \frac{d\mathbb{Q}}{d\mathbb{P}} \right|_{\mathcal{F}_t} = \sum_{v \in N(t)} e^{-\int_0^t M(X_v(s))R(X_v(s))ds} \zeta_v(t) =: Z(t)$$

and Z(t) is a \mathbb{P} -martingale with respect to \mathcal{F}_t (again see Theorem 2.4 of [3] for details). Since Z(t) is a positive martingale, it converges \mathbb{P} -almost surely to $Z(\infty) := \liminf Z(t)$.

We now state our main result. Suppose that f(t) is $\tilde{\mathcal{F}}_t$ -measurable for each t. Then, again, we may write each f(t) via the representation

$$f(t) = \sum_{u \in N_t} f_u(t) \mathbb{1}_{\{\text{spine}(t) = u\}}$$

where $f_u(t)$ is \mathcal{F}_t -measurable for each $t \ge 0$ and each $u \in N(t)$.

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Theorem 1.1. Suppose that $\{f(t) : t \ge 0\}$ is $\tilde{\mathbb{Q}}$ -uniformly integrable. If $f(t) \to f \tilde{\mathbb{Q}}$ almost surely as $t \to \infty$ then

$$\sum_{u \in N_t} f_u(t) \frac{e^{-\int_0^t M(X_u(s))R(X_u(s))ds} \zeta_u(t)}{Z(t)} \to \tilde{\mathbb{Q}}[f|\mathcal{F}_\infty] \tag{(*)}$$

Q-almost surely. Furthermore, $\mathbb{P}((\star)|Z(\infty) > 0) = 1$.

Remark 1.2. 1. Note that

$$\frac{e^{-\int_0^t M(X_u(s))R(X_u(s))ds}\zeta_u(t)}{Z(t)} = \tilde{\mathbb{Q}}(\text{spine}(t) = u|\mathcal{F}_t).$$

2. Since 1/Z(t) is a positive Q-supermartingale (and thus converges almost surely to an almost surely finite limit), $Z(t) \rightarrow Z(\infty)$ Q-almost surely. Thus we may deduce from (*) that

$$\sum_{u \in N_t} f_u(t) e^{-\int_0^t M(X_u(s))R(X_u(s))ds} \zeta_u(t) \to \tilde{\mathbb{Q}}[f|\mathcal{F}_\infty]Z(\infty) \quad \mathbb{Q}\text{-almost surely.}$$

In fact under fairly mild conditions on the branching distributions A(x), we have $\mathbb{Q}(Z(\infty) < \infty) = 1$, in which case we do not lose anything by rewriting (*) in this way.

3. In many cases of interest the events $\{Z(\infty) = 0\}$ and $\{\exists t \in [0, \infty) : Z(t) = 0\}$ agree to within a set of zero \mathbb{P} -probability. Then, of course,

$$\sum_{u \in N_t} f_u(t) e^{-\int_0^t M(X_u(s))R(X_u(s))ds} \zeta_u(t) \to \tilde{\mathbb{Q}}[f|\mathcal{F}_\infty]Z(\infty) \quad \mathbb{P}\text{-almost surely.}$$

2 Some example applications

We outline here two examples showing how our strong law can be applied. The first example is folklore in branching processes, but we are not aware of another proof.

In branching processes branching at rate β into on average *m* offspring, most particles branch at rate $m\beta$.

Take a continuous-time branching process with constant birth rate $R(x) \equiv \beta$ and birth distribution $A(x) \equiv A$ satisfying $\mathbb{E}[A \log_+ A] < \infty$ with $m := \mathbb{E}[A]$. Let $\zeta(t) \equiv 1$. For any $\varepsilon > 0$ we may take $f(t) = \mathbb{1}_{\{|n_t/t - m\beta| < \varepsilon\}}$, the indicator that birth rate along the spine up to time t is close to its expected value under $\tilde{\mathbb{Q}}$, $m\beta$. Then for any $\varepsilon > 0$, f(t) converges $\tilde{\mathbb{Q}}$ -almost surely to 1. Thus Theorem 1.1, together with some classical results on branching processes concerning the martingale $e^{-(m-1)\beta t}|N(t)|$, tells us that on the event that the process survives,

$$\frac{1}{|N(t)|} \sum_{u \in N(t)} f_u(t) \to 1 \quad \mathbb{P}\text{-almost surely.}$$

This may be interpreted as saying that if we choose a particle uniformly at random from those alive at a large time t, and look at its history, we are likely to see that its average birth rate has been approximately $m\beta$. In particular, with binary branching, we see an average birth rate of 2β in typical particles (rather than β , which one might naively expect).

Our second example shows how the spatial behaviour of the spine can also be passed to other particles: if the spine shows ergodic behaviour, then so do many other particles.

Occupation densities and ergodic spines.

Suppose that the motion of the spine $(\xi_t, t \ge 0)$ is ergodic under \mathbb{Q} with invariant probability density π in the sense that there exists some suitable class of functions \mathcal{H} such that for any $h \in \mathcal{H}$,

$$\frac{1}{t} \int_0^t h(\xi_s) ds \to L_h := \int_{\mathbb{R}} h(x) \pi(x) dx \quad \tilde{\mathbb{Q}}\text{-almost surely.}$$

Then for any continuous function $g : \mathbb{R} \to \mathbb{R}$ and any $h \in \mathcal{H}$,

$$\frac{1}{Z(t)}\sum_{u\in N(t)}g\left(\frac{1}{t}\int_0^t h(X_u(s))ds\right)e^{-\int_0^t M(X_u(s))R(X_u(s))ds}\zeta_u(t)\to g(L_h)$$

 \mathbb{Q} -almost surely. The same holds under \mathbb{P} on the event $Z(\infty) > 0$, which is one exposition of the general principle that if forcing the spine to show certain behaviour does not cause the corresponding martingale to disappear, then that behaviour appears in the original process.

Particles within a tube.

To give a more explicit example using the same idea as above, we consider a standard branching Brownian motion in which all particles breed at rate 1 into two new particles. That is, ψ_t is a Brownian motion, $R \equiv 1$ and $A \equiv 2$ so $M \equiv 1$. We begin with one particle at $x \in (0, L)$, and we are interested in particles that remain between 0 and L. Introduce the martingale

$$\zeta(t) = a e^{\pi^2 t/2L^2} \sin\left(\frac{\pi\xi_t}{L}\right) \mathbb{1}_{\{\xi_s \in (0,L) \ \forall s \le t\}}$$

where $a = 1/\sin(\pi x/L)$, which gives

$$Z(t) = a \sum_{u \in N(t)} e^{-t + \pi^2 t/2L^2} \sin\left(\frac{\pi X_u(t)}{L}\right) \mathbb{1}_{\{X_u(s) \in (0,L) \ \forall s \le t\}}.$$

The invariant density of the motion of the spine under $\tilde{\mathbb{Q}}$ is $\frac{2}{L}\sin^2(\pi y/L)$, $y \in (0, L)$, so for any Borel set A and any $\varepsilon > 0$,

$$\mathbb{1}_{\left\{\left|\frac{1}{t}\int_{0}^{t}\mathbb{1}_{\left\{\xi_{s}\in A\right\}}ds-\frac{2}{L}\int_{A}\sin^{2}(\pi y/L)dy\right|<\varepsilon\right\}}\to 1 \quad \mathbb{Q}-\text{almost surely}$$

Thus, setting

$$\zeta_u(t) = a e^{\pi^2 t/2L^2} \sin\left(\frac{\pi X_u(t)}{L}\right) \mathbb{1}_{\{X_u(s) \in (0,L) \, \forall s \le t\}}$$

we have that

$$\frac{1}{Z(t)} \sum_{u \in N(t)} e^{-t} \zeta_u(t) \mathbb{1}_{\left\{ \left| \frac{1}{t} \int_0^t \mathbb{1}_{\left\{ X_u(s) \in A \right\}} ds - \frac{2}{L} \int_A \sin^2(\pi y/L) dy \right| < \varepsilon \right\}} \to 1$$

Q-almost surely and, when $L > \pi/\sqrt{2}$, P-almost surely. This result was applied to prove Lemma 5 of [2] — see also [2] for a proof that $Z(\infty) > 0$ when $L > \pi/\sqrt{2}$.

3 Measure theoretic results

To prove Theorem 1.1 we need some simple measure theory. For this section we forget the branching setup and take any filtered probability space $(\Omega, \mathcal{F}, \mathcal{F}_t, P)$ and define $\mathcal{F}_{\infty} := \bigvee_{t \geq 0} \mathcal{F}_t$. Suppose that $X_t, t \geq 0$ is an \mathcal{F} -measurable process such that $(E[X_t|\mathcal{F}_t], t \geq 0)$ is almost surely càdlàg.

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Proposition 3.1. If

$$E[X_t | \mathcal{F}_{\infty}] \rightarrow Y$$
 almost surely,

then

$$E[X_t | \mathcal{F}_t] \to Y$$
 almost surely.

Proof. Fix $\varepsilon > 0$. We show that there exists an almost surely finite random variable T such that

$$\sup_{t>T} E[X_t | \mathcal{F}_t] \le Y + \varepsilon \text{ almost surely.}$$

By the càdlàg property, it is sufficient to take the supremum above over rationals greater than T; from now on all our suprema will be over rationals.

Since $E[X_t|\mathcal{F}_{\infty}] \to Y$, there exists an almost surely finite random variable T_1 such that

$$\sup_{t \geq T_1} E[X_t | \mathcal{F}_\infty] < Y + \varepsilon/2 \; \text{ almost surely,}$$

and since (by the fact that it is a closed martingale) $E[Y|\mathcal{F}_t] \to E[Y|\mathcal{F}_\infty] = Y$ (Y is \mathcal{F}_∞ -measurable since it is the limit of \mathcal{F}_∞ -measurable random variables), there exists an almost surely finite random variable T_2 such that

$$\sup_{t \ge T_2} E[Y|\mathcal{F}_t] < Y + \varepsilon/2 \text{ almost surely.}$$

Let $T = T_1 \vee T_2$. Then

$$\sup_{t \ge T} E[X_t | \mathcal{F}_t] \leq \sup_{t \ge T_2} \sup_{s \ge T_1} E[X_s | \mathcal{F}_t]$$

$$= \sup_{t \ge T_2} \sup_{s \ge T_1} E[E[X_s | \mathcal{F}_\infty] | \mathcal{F}_t]$$

$$\leq \sup_{t \ge T_2} E\left[\sup_{s \ge T_1} E[X_s | \mathcal{F}_\infty] \right| \mathcal{F}_t \right]$$

$$\leq \sup_{t \ge T_2} E[Y + \varepsilon/2 | \mathcal{F}_t]$$

$$\leq Y + \varepsilon$$

(all statements hold almost surely). Thus $\limsup E[X_t | \mathcal{F}_t] \leq Y$; the proof that $\liminf E[X_t | \mathcal{F}_t] \geq Y$ is similar. \Box

Corollary 3.2. Suppose that the collection of random variables $\{X_t, t \ge 0\}$ is uniformly integrable. If

$$X_t \to X$$
 almost surely

then

$$E[X_t | \mathcal{F}_t] o E[X | \mathcal{F}_\infty]$$
 almost surely.

Proof. Let $Y = E[X|\mathcal{F}_{\infty}]$; then by uniform integrability,

$$E[X_t | \mathcal{F}_{\infty}] \to Y$$
 almost surely.

Proposition 1 now gives the result.

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4 The proof of Theorem 1.1

We now return to the notation from Section 1.

Proof of Theorem 1.1. We recall Theorem 8.2 of Hardy and Harris [1], which says that under the conditions above,

$$\tilde{\mathbb{Q}}[f(t)|\mathcal{F}_t] = \sum_{u \in N_t} f_u(t) \frac{e^{-\int_0^t M(X_u(s))R(X_u(s))ds} \zeta_u(t)}{Z(t)}.$$

Now if f(t) converges $\tilde{\mathbf{Q}}\text{-almost surely to }f$ then by Corollary 3.2 we have

 $\tilde{\mathbb{Q}}[f(t)|\mathcal{F}_t] \to \tilde{\mathbb{Q}}[f|\mathcal{F}_\infty] \ \tilde{\mathbb{Q}}$ -almost surely

and hence

$$\sum_{u \in N_t} f_u(t) \frac{e^{-\int_0^t M(X_u(s))R(X_u(s))ds} \zeta_u(t)}{Z(t)} \to \tilde{\mathbb{Q}}[f|\mathcal{F}_\infty] \ \tilde{\mathbb{Q}}\text{-almost surely.}$$

Finally, for any (\mathcal{F}_{∞} -measurable) event A such that $\mathbb{Q}(A) = 1$,

$$\mathbb{P}(A|Z(\infty) > 0) = \frac{\mathbb{P}(A \cap \{Z(\infty) > 0\})}{\mathbb{P}(Z(\infty) > 0)}$$
$$= \frac{\mathbb{Q}\left[\frac{1}{Z(\infty)}\mathbb{1}_{A \cap \{Z(\infty) > 0\}}\right]}{\mathbb{P}(Z(\infty) > 0)}$$
$$= \frac{\mathbb{Q}\left[\frac{1}{Z(\infty)}\mathbb{1}_{\{Z(\infty) > 0\}}\right]}{\mathbb{P}(Z(\infty) > 0)}$$
$$= 1.$$

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