Research Article

Adaptive Wavelet Estimation of a Biased Density for Strongly Mixing Sequences

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The estimation of a biased density for exponentially strongly mixing sequences is investigated. We construct a new adaptive wavelet estimator based on a hard thresholding rule. We determine a sharp upper bound of the associated mean integrated square error for a wide class of functions.

1. Introduction

In the standard density estimation problem, we observe *n* random variables $X_1, ..., X_n$ with common density function *f*. The goal is to estimate *f* from $X_1, ..., X_n$. However, in some applications, $X_1, ..., X_n$ are not accessible; we only have *n* random variables $Z_1, ..., Z_n$ with the common density

$$g(x) = \mu^{-1} w(x) f(x), \qquad (1.1)$$

where *w* denotes a known positive function and μ is the unknown normalization parameter: $\mu = \int w(y) f(y) dy$. Our goal is to estimate the "biased density" *f* from Z_1, \ldots, Z_n . Practical examples can be found in, for example, [1–3] and the survey by the author of [4].

The standard i.i.d. case has been investigated in several papers. See, for example, [5–9]. To the best of our knowledge, the dependent case has only been examined in [10] for associated (positively or negatively) Z_1, \ldots, Z_n . In this paper, we study another dependent (and realistic) structure which has not been addressed earlier: we suppose that Z_1, \ldots, Z_n is a sample of a strictly stationary and exponentially strongly mixing process $(Z_i)_{i \in \mathbb{Z}}$ (to be defined in Section 2). Such a dependence condition arises for a wide class of GARCH-type time series models classically encountered in finance. See, for example, [11, 12] for an overview.

We focus our attention on the wavelet methods because they provide a coherent set of procedures that are spatially adaptive and near optimal over a wide range of function spaces. See, for example, [13, 14] for a detailed coverage of wavelet theory in statistics. We develop two new wavelet estimators: a linear nonadaptive based on projections and a nonlinear adaptive using the hard thresholding rule introduced by [15]. We measure their performances by determining upper bounds of the mean integrated squared error (MISE) over Besov balls (to be defined in Section 3). We prove that our adaptive estimator attains a sharp rate of convergence, close to the one attained by the linear wavelet estimator (constructed in a nonadaptive fashion to minimize the MISE).

The rest of the paper is organized as follows. Section 2 is devoted to the assumptions on the model. In Section 3, we present wavelets and Besov balls. The considered wavelet estimators are defined in Section 4. Section 5 is devoted to the results. The proofs are postponed in Section 6.

2. Assumptions on the Model

We assume that $Z_1, ..., Z_n$ coming from a strictly stationary process $(Z_i)_{i \in \mathbb{Z}}$. For any $m \in \mathbb{Z}$, we define the *m*th strongly mixing coefficient of $(Z_i)_{i \in \mathbb{Z}}$ by

$$a_m = \sup_{(A,B)\in\mathcal{F}^Z_{-\infty,0}\times\mathcal{F}^Z_{m,\infty}} |\mathbb{P}(A\cap B) - \mathbb{P}(A)\mathbb{P}(B)|,$$
(2.1)

where, for any $u \in \mathbb{Z}$, $\mathcal{F}_{-\infty,u}^Z$ is the σ -algebra generated by the random variables ..., Z_{u-1} , Z_u and $\mathcal{F}_{u,\infty}^Z$ is the σ -algebra generated by the random variables Z_u, Z_{u+1}, \ldots

We consider the exponentially strongly mixing case, that is, there exist three known constants, $\gamma > 0$, c > 0, and $\theta > 0$, such that, for any $m \in \mathbb{Z}$,

$$a_m \le \gamma \exp\left(-c|m|^{\theta}\right). \tag{2.2}$$

This assumption is satisfied by a large class of GARCH processes. See, for example, [11, 12, 16, 17].

Note that, when $\theta \to \infty$, we are in the standard i.i.d. case. W.o.l.g., the support of the functions f, and w are [0,1]. There exist two constants, c > 0 and C > 0, such that

$$c \le \inf_{x \in [0,1]} w(x), \qquad \sup_{x \in [0,1]} w(x) \le C.$$
 (2.3)

There exists a (known) constant C > 0 such that

$$\sup_{x \in [0,1]} f(x) \le C.$$
(2.4)

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For any $m \in \{1, ..., n\}$, let $g_{(Z_0, Z_m)}$ be the density of (Z_0, Z_m) . There exists a constant C > 0 such that

$$\sup_{m \in \{1,\dots,n\}} \sup_{(x,y) \in [0,1]^2} |g_{(Z_0,Z_m)}(x,y) - g(x)g(y)| \le C.$$
(2.5)

The two first boundedness assumptions are standard in the estimation of biased densities. See, for example, [6–8].

3. Wavelets and Besov Balls

Let *N* be an integer ϕ and ψ be the initial wavelets of dbN (so supp(ϕ) = supp(ψ) = [1 – *N*, *N*]). Set

$$\phi_{j,k}(x) = 2^{j/2} \phi \Big(2^j x - k \Big), \qquad \psi_{j,k}(x) = 2^{j/2} \psi \Big(2^j x - k \Big). \tag{3.1}$$

With an appropriate treatments at the boundaries, there exists an integer τ satisfying $2^{\tau} \ge 2N$ such that the collection $\mathcal{B} = \{\phi_{\tau,k}(\cdot), k \in \{0, ..., 2^{\tau} - 1\}; \psi_{j,k}(\cdot); j \in \mathbb{N} - \{0, ..., \tau - 1\}, k \in \{0, ..., 2^{j} - 1\}\}$, is an orthonormal basis of $\mathbb{L}^{2}([0, 1])$ (the space of square-integrable functions on [0, 1]). See [18].

For any integer $\ell \ge \tau$, any $h \in \mathbb{L}^2([0,1])$ can be expanded on \mathcal{B} as

$$h(x) = \sum_{k=0}^{2^{\ell}-1} \alpha_{\ell,k} \phi_{\ell,k}(x) + \sum_{j=\ell}^{\infty} \sum_{k=0}^{2^{j}-1} \beta_{j,k} \psi_{j,k}(x), \quad x \in [0,1],$$
(3.2)

where $\alpha_{j,k}$ and $\beta_{j,k}$ are the wavelet coefficients of *h* defined by

$$\alpha_{j,k} = \int_0^1 h(x)\phi_{j,k}(x)dx, \qquad \beta_{j,k} = \int_0^1 h(x)\psi_{j,k}(x)dx.$$
(3.3)

Let M > 0, s > 0, $p \ge 1$, and $r \ge 1$. A function h belongs to $B_{p,r}^s(M)$ if and only if there exists a constant $M^* > 0$ (depending on M) such that the associated wavelet coefficients (3.3) satisfy

$$2^{\tau(1/2-1/p)} \left(\sum_{k=0}^{2^{\tau}-1} |\alpha_{\tau,k}|^p\right)^{1/p} + \left(\sum_{j=\tau}^{\infty} \left(2^{j(s+1/2-1/p)} \left(\sum_{k=0}^{2^{j}-1} |\beta_{j,k}|^p\right)^{1/p}\right)^r\right)^{1/r} \le M^*.$$
(3.4)

In this expression, *s* is a smoothness parameter and *p* and *r* are norm parameters. For a particular choice of *s*, *p*, and *r*, $B_{p,r}^s(M)$ contains some classical sets of functions as the Hölder and Sobolev balls. See [19].

4. Estimators

Firstly, we consider the following estimator for μ :

$$\widehat{\mu} = \left(\frac{1}{n} \sum_{i=1}^{n} \frac{1}{w(Z_i)}\right)^{-1}.$$
(4.1)

It is obtained by the method of moments (see Proposition 6.2 below).

Then, for any integer $j \ge \tau$ and any $k \in \{0, ..., 2^j - 1\}$, we estimate the unknown wavelet coefficient

(i)
$$\alpha_{j,k} = \int_0^1 f(x)\phi_{j,k}(x)dx$$
 by

$$\hat{\alpha}_{j,k} = \frac{\hat{\mu}}{n} \sum_{i=1}^n \frac{\phi_{j,k}(Z_i)}{w(Z_i)},$$
(4.2)

(ii)
$$\beta_{j,k} = \int_0^1 f(x) \varphi_{j,k}(x) dx$$
 by

$$\widehat{\beta}_{j,k} = \frac{\widehat{\mu}}{n} \sum_{i=1}^{n} \frac{\psi_{j,k}(Z_i)}{w(Z_i)}.$$
(4.3)

Note that they are those considered in the i.i.d. case (see, e.g., [8, 9]). Their statistical properties, with our dependent structure, are investigated in Propositions 6.2, 6.3, and 6.4 below.

Assuming that $f \in B^s_{p,r}(M)$ with $p \ge 2$, we define the linear estimator \widehat{f}^L by

$$\widehat{f}^{L}(x) = \sum_{k=0}^{2^{j_0}-1} \widehat{\alpha}_{j_0,k} \phi_{j_0,k}(x), \quad x \in [0,1],$$
(4.4)

where $\hat{\alpha}_{j,k}$ is defined by (4.2) and j_0 is the integer satisfying

$$\frac{1}{2}n^{1/(2s+1)} < 2^{j_0} \le n^{1/(2s+1)}.$$
(4.5)

For a survey on wavelet linear estimators for various density models, we refer the reader to [20]. For the consideration of strongly mixing sequences, see, for example, [21, 22]. We define the hard thresholding estimator \hat{f}^H by

$$\widehat{f}^{H}(x) = \sum_{k=0}^{2^{\tau}-1} \widehat{\alpha}_{\tau,k} \phi_{\tau,k}(x) + \sum_{j=\tau}^{j_{1}} \sum_{k=0}^{2^{j}-1} \widehat{\beta}_{j,k} \mathbb{I}_{\{|\widehat{\beta}_{j,k}| \ge \kappa \lambda_{n}\}} \psi_{j,k}(x),$$
(4.6)

 $x \in [0,1]$, where $\hat{\alpha}_{\tau,k}$ is defined by (4.2) and $\hat{\beta}_{j,k}$ by (4.3), for any random event \mathcal{A} , $\mathbb{I}_{\mathcal{A}}$ is the indicator function on \mathcal{A} , j_1 is the integer satisfying

$$\frac{1}{2} \frac{n}{\left(\ln n\right)^{1+1/\theta}} < 2^{j_1} \le \frac{n}{\left(\ln n\right)^{1+1/\theta}},\tag{4.7}$$

 θ is the one in (2.2), κ is a large enough constant (the one in Proposition 6.4 below) and λ_n is the threshold

$$\lambda_n = \sqrt{\frac{(\ln n)^{1+1/\theta}}{n}}.$$
(4.8)

The feature of the hard thresholding estimator is to only estimate the "large" unknown wavelet coefficients of f which contain his main characteristics.

For the construction of hard thresholding wavelet estimators in the standard density model, see, for example, [15, 23].

5. Results

Theorem 5.1 (upper bound for \hat{f}^L). Consider (1.1) under the assumptions of Section 2. Suppose that $f \in B^s_{p,r}(M)$ with s > 0, $p \ge 2$, and $r \ge 1$. Let \hat{f}^L be (4.4). Then there exists a constant C > 0 such that

$$\mathbb{E}\left(\int_0^1 \left(\widehat{f}^L(x) - f(x)\right)^2 dx\right) \le C n^{-2s/(2s+1)}.$$
(5.1)

The proof of Theorem 5.1 uses a suitable decomposition of the MISE and a moment inequality on (4.2) (see Proposition 6.3 below).

Note that $n^{-2s/(2s+1)}$ is the optimal rate of convergence (in the minimax sense) for the standard density model in the independent case (see, e.g., [14, 23]).

Theorem 5.2 (upper bound for \hat{f}^H). Consider (1.1) under the assumptions of Section 2. Let \hat{f}^H be (4.6). Suppose that $f \in B^s_{p,r}(M)$ with $r \ge 1$, $\{p \ge 2 \text{ and } s > 0\}$ or $\{p \in [1,2) \text{ and } s > 1/p\}$. Then there exists a constant C > 0 such that

$$\mathbb{E}\left(\int_0^1 \left(\widehat{f}^H(x) - f(x)\right)^2 dx\right) \le C\left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2s/(2s+1)}.$$
(5.2)

The proof of Theorem 5.2 uses a suitable decomposition of the MISE, some moment inequalities on (4.2) and (4.3) (see Proposition 6.3 below), and a concentration inequality on (4.3) (see Proposition 6.4 below).

Theorem 5.2 shows that, besides being adaptive, \hat{f}^H attains a rate of convergence close to the one of \hat{f}^L . The only difference is the logarithmic term $(\ln n)^{(1+1/\theta)(2s/(2s+1))}$.

Note that, if we restrict our study to the independent case, that is, $\theta \to \infty$, the rate of convergence attained by \hat{f}^H becomes the standard one: $(\log n/n)^{2s/(2s+1)}$. See, for example, [14, 15, 23].

6. Proofs

In this section, we consider (1.1) under the assumptions of Section 2. Moreover, *C* denotes any constant that does not depend on *j*, *k* and *n*. Its value may change from one term to another and may depends on ϕ or ψ .

6.1. Auxiliary Results

Lemma 6.1. For any integer $j \ge \tau$ and any $k \in \{0, ..., 2^j - 1\}$, let $\hat{\alpha}_{j,k}$ be (4.2) and $\alpha_{j,k} = \int_0^1 f(x)\phi_{j,k}(x)dx$. Then, under the assumptions of Section 2, there exists a constant C > 0 such that

$$\left|\widehat{\alpha}_{j,k} - \alpha_{j,k}\right| \le C\left(\left|\frac{\mu}{n}\sum_{i=1}^{n}\frac{\phi_{j,k}(Z_i)}{w(Z_i)} - \alpha_{j,k}\right| + \left|\frac{1}{\widehat{\mu}} - \frac{1}{\mu}\right|\right).$$
(6.1)

This inequality holds for ψ instead of ϕ (and, a fortiori, $\hat{\beta}_{j,k}$ defined by (4.3) instead of $\hat{\alpha}_{j,k}$ and $\beta_{j,k} = \int_0^1 f(x)\psi_{j,k}(x)dx$ instead of $\alpha_{j,k}$).

Proof of Lemma 6.1. We have

$$\widehat{\alpha}_{j,k} - \alpha_{j,k} = \frac{\widehat{\mu}}{\mu} \left(\frac{\mu}{n} \sum_{i=1}^{n} \frac{\phi_{j,k}(Z_i)}{w(Z_i)} - \alpha_{j,k} \right) + \alpha_{j,k} \widehat{\mu} \left(\frac{1}{\mu} - \frac{1}{\widehat{\mu}} \right).$$
(6.2)

Due to (2.3), we have $|\hat{\mu}| \leq C$ and $|\hat{\mu}/\mu| \leq C$. Therefore

$$\left|\widehat{\alpha}_{j,k} - \alpha_{j,k}\right| \le C\left(\left|\frac{\mu}{n}\sum_{i=1}^{n}\frac{\phi_{j,k}(Z_i)}{w(Z_i)} - \alpha_{j,k}\right| + \left|\alpha_{j,k}\right| \left|\frac{1}{\widehat{\mu}} - \frac{1}{\mu}\right|\right).$$
(6.3)

Using (2.4) and the Cauchy-Schwarz inequality, we obtain

$$|\alpha_{j,k}| \leq \int_{0}^{1} f(x) |\phi_{j,k}(x)| dx \leq C \int_{0}^{1} |\phi_{j,k}(x)| dx$$

$$\leq C \left(\int_{0}^{1} (\phi_{j,k}(x))^{2} dx \right)^{1/2} = C.$$
(6.4)

Hence

$$\left|\widehat{\alpha}_{j,k} - \alpha_{j,k}\right| \le C\left(\left|\frac{\mu}{n}\sum_{i=1}^{n}\frac{\phi_{j,k}(Z_i)}{w(Z_i)} - \alpha_{j,k}\right| + \left|\frac{1}{\widehat{\mu}} - \frac{1}{\mu}\right|\right).$$
(6.5)

Lemma 6.1 is proved.

Proposition 6.2. For any integer $j \ge \tau$ such that $2^j \le n$ and any $k \in \{0, \ldots, 2^j - 1\}$, let $\alpha_{j,k} = \int_0^1 f(x)\phi_{j,k}(x)dx$ and $\hat{\mu}$ be (4.1). Then,

(1) *one has*

$$\mathbb{E}\left(\frac{\mu}{n}\sum_{i=1}^{n}\frac{\phi_{j,k}(Z_i)}{w(Z_i)}\right) = \alpha_{j,k}, \qquad \mathbb{E}\left(\frac{1}{\widehat{\mu}}\right) = \frac{1}{\mu}, \tag{6.6}$$

(2) there exists a constant C > 0 such that

$$\mathbb{V}\left(\frac{\mu}{n}\sum_{i=1}^{n}\frac{\phi_{j,k}(Z_i)}{w(Z_i)}\right) \le C\frac{1}{n},\tag{6.7}$$

(3) there exists a constant C > 0 such that

$$\mathbb{V}\left(\frac{1}{\hat{\mu}}\right) \le C\frac{1}{n}.\tag{6.8}$$

These results hold for ψ instead of ϕ (and, a fortiori, $\beta_{j,k} = \int_0^1 f(x)\psi_{j,k}(x)dx$ instead of $\alpha_{j,k}$). Proof of Proposition 6.2. (1) We have

$$\mathbb{E}\left(\frac{\mu}{n}\sum_{i=1}^{n}\frac{\phi_{j,k}(Z_{i})}{w(Z_{i})}\right) = \mu\mathbb{E}\left(\frac{\phi_{j,k}(Z_{1})}{w(Z_{1})}\right) = \mu\int_{0}^{1}\frac{\phi_{j,k}(x)}{w(x)}g(x)dx$$

$$= \mu\int_{0}^{1}\frac{\phi_{j,k}(x)}{w(x)}\mu^{-1}w(x)f(x)dx = \int_{0}^{1}f(x)\phi_{j,k}(x)dx = \alpha_{j,k}.$$
(6.9)

Since f is a density, we obtain

$$\mathbb{E}\left(\frac{1}{\hat{\mu}}\right) = \mathbb{E}\left(\frac{1}{n}\sum_{i=1}^{n}\frac{1}{w(Z_{i})}\right) = \mathbb{E}\left(\frac{1}{w(Z_{1})}\right) = \int_{0}^{1}\frac{1}{w(x)}g(x)dx$$

$$= \int_{0}^{1}\frac{1}{w(x)}\mu^{-1}w(x)f(x)dx = \frac{1}{\mu}\int_{0}^{1}f(x)dx = \frac{1}{\mu}.$$
(6.10)

(2) We have

$$\mathbb{V}\left(\frac{\mu}{n}\sum_{i=1}^{n}\frac{\phi_{j,k}(Z_{i})}{w(Z_{i})}\right) = \frac{\mu^{2}}{n^{2}}\sum_{v=1}^{n}\sum_{\ell=1}^{n}\mathbb{C}\left(\frac{\phi_{j,k}(Z_{v})}{w(Z_{v})}, \frac{\phi_{j,k}(Z_{\ell})}{w(Z_{\ell})}\right) \\
= \frac{\mu^{2}}{n}\mathbb{V}\left(\frac{\phi_{j,k}(Z_{1})}{w(Z_{1})}\right) + 2\frac{\mu^{2}}{n^{2}}\sum_{v=2}^{n}\sum_{\ell=1}^{v-1}\mathbb{C}\left(\frac{\phi_{j,k}(Z_{v})}{w(Z_{v})}, \frac{\phi_{j,k}(Z_{\ell})}{w(Z_{\ell})}\right) \\
\leq \frac{\mu^{2}}{n}\mathbb{V}\left(\frac{\phi_{j,k}(Z_{1})}{w(Z_{1})}\right) + 2\frac{\mu^{2}}{n^{2}}\left|\sum_{v=2}^{n}\sum_{\ell=1}^{v-1}\mathbb{C}\left(\frac{\phi_{j,k}(Z_{v})}{w(Z_{v})}, \frac{\phi_{j,k}(Z_{\ell})}{w(Z_{\ell})}\right)\right|.$$
(6.11)

Using (2.3) and (2.4), we have $\sup_{x \in [0,1]} g(x) \le C$. Hence,

$$\mathbb{V}\left(\frac{\phi_{j,k}(Z_1)}{w(Z_1)}\right) \leq \mathbb{E}\left(\left(\frac{\phi_{j,k}(Z_1)}{w(Z_1)}\right)^2\right) \leq C\mathbb{E}\left(\left(\phi_{j,k}(Z_1)\right)^2\right)$$

$$= C\int_0^1 \left(\phi_{j,k}(x)\right)^2 g(x) dx \leq C\int_0^1 \left(\phi_{j,k}(x)\right)^2 dx = C.$$
(6.12)

It follows from the stationarity of $(Z_i)_{i \in \mathbb{Z}}$ and $2^j \leq n$ that

$$\left| \sum_{v=2}^{n} \sum_{\ell=1}^{v-1} \mathbb{C}\left(\frac{\phi_{j,k}(Z_v)}{w(Z_v)}, \frac{\phi_{j,k}(Z_\ell)}{w(Z_\ell)}\right) \right| = \left| \sum_{m=1}^{n} (n-m) \mathbb{C}\left(\frac{\phi_{j,k}(Z_0)}{w(Z_0)}, \frac{\phi_{j,k}(Z_m)}{w(Z_m)}\right) \right|$$

$$\leq n \sum_{m=1}^{n} \left| \mathbb{C}\left(\frac{\phi_{j,k}(Z_0)}{w(Z_0)}, \frac{\phi_{j,k}(Z_m)}{w(Z_m)}\right) \right| = T_1 + T_2,$$
(6.13)

where

$$T_{1} = n \sum_{m=1}^{2^{j-1}} \left| \mathbb{C}\left(\frac{\phi_{j,k}(Z_{0})}{w(Z_{0})}, \frac{\phi_{j,k}(Z_{m})}{w(Z_{m})}\right) \right|,$$

$$T_{2} = n \sum_{m=2^{j}}^{n} \left| \mathbb{C}\left(\frac{\phi_{j,k}(Z_{0})}{w(Z_{0})}, \frac{\phi_{j,k}(Z_{m})}{w(Z_{m})}\right) \right|.$$
(6.14)

Let us now bound T_1 and T_2 .

Upper Bound for T_1

Using (2.5), (2.3), and doing the change a variables $y = 2^{j}x - k$, we obtain

$$\left| \mathbb{C} \left(\frac{\phi_{j,k}(Z_0)}{w(Z_0)}, \frac{\phi_{j,k}(Z_m)}{w(Z_m)} \right) \right| = \left| \iint_0^1 (g_{(Z_0, Z_m)}(x, y) - g(x)g(y)) \frac{\phi_{j,k}(x)}{w(x)} \frac{\phi_{j,k}(y)}{w(y)} dx \, dy \right|$$

$$\leq \iint_0^1 |g_{(Z_0, Z_m)}(x, y) - g(x)g(y)| \left| \frac{\phi_{j,k}(x)}{w(x)} \right| \left| \frac{\phi_{j,k}(y)}{w(y)} \right| dx \, dy$$

$$\leq C \left(\int_0^1 |\phi_{j,k}(x)| dx \right)^2 = C \left(2^{-j/2} \int_0^1 |\phi(x)| dx \right)^2 = C 2^{-j}.$$

(6.15)

Therefore,

$$T_1 \le Cn2^{-j}2^j = Cn. (6.16)$$

Upper Bound for T_2

By the Davydov inequality for strongly mixing processes (see [24]), for any $q \in (0, 1)$, it holds that

$$\left| \mathbb{C}\left(\frac{\phi_{j,k}(Z_0)}{w(Z_0)}, \frac{\phi_{j,k}(Z_m)}{w(Z_m)}\right) \right| \le 10a_m^q \left(\mathbb{E}\left(\left|\frac{\phi_{j,k}(Z_0)}{w(Z_0)}\right|^{2/(1-q)}\right) \right)^{1-q}$$

$$\le 10a_m^q \left(\sup_{x \in [0,1]} \left|\frac{\phi_{j,k}(x)}{w(x)}\right| \right)^{2q} \left(\mathbb{E}\left(\left(\frac{\phi_{j,k}(Z_0)}{w(Z_0)}\right)^2 \right) \right)^{1-q}.$$

$$(6.17)$$

By (2.3), we have

$$\sup_{x \in [0,1]} \left| \frac{\phi_{j,k}(x)}{w(x)} \right| \le C \sup_{x \in [0,1]} \left| \phi_{j,k}(x) \right| \le C 2^{j/2}$$
(6.18)

and, by (6.12),

$$\mathbb{E}\left(\left(\frac{\phi_{j,k}(Z_0)}{w(Z_0)}\right)^2\right) \le C.$$
(6.19)

Therefore,

$$\left| \mathbb{C}\left(\frac{\phi_{j,k}(Z_0)}{w(Z_0)}, \frac{\phi_{j,k}(Z_m)}{w(Z_m)}\right) \right| \le C 2^{qj} a_m^q.$$
(6.20)

Since $\sum_{m=2^j}^n m^q a_m^q \leq \sum_{m=1}^\infty m^q a_m^q = \gamma^q \sum_{m=1}^\infty m^q \exp(-cqm^\theta) < \infty$, we have

$$T_2 \le Cn2^{qj} \sum_{m=2^j}^n a_m^q \le Cn \sum_{m=2^j}^n m^q a_m^q \le Cn.$$
(6.21)

It follows from (6.13), (6.16), and (6.21) that

$$\left|\sum_{v=2}^{n}\sum_{\ell=1}^{v-1}\mathbb{C}\left(\frac{\phi_{j,k}(Z_v)}{w(Z_v)},\frac{\phi_{j,k}(Z_\ell)}{w(Z_\ell)}\right)\right| \le Cn.$$
(6.22)

Combining (6.11), (6.12), and (6.22), we obtain

$$\mathbb{V}\left(\frac{\mu}{n}\sum_{i=1}^{n}\frac{\phi_{j,k}(Z_i)}{w(Z_i)}\right) \le C\frac{1}{n}.$$
(6.23)

(3) Proceeding in a similar fashion to 2-, we obtain

$$\mathbb{V}\left(\frac{1}{\hat{\mu}}\right) = \mathbb{V}\left(\frac{1}{n}\sum_{i=1}^{n}\frac{1}{w(Z_{i})}\right)$$
$$= \frac{1}{n}\mathbb{V}\left(\frac{1}{w(Z_{1})}\right) + 2\frac{1}{n^{2}}\sum_{v=2}^{n}\sum_{\ell=1}^{v-1}\mathbb{C}\left(\frac{1}{w(Z_{v})}, \frac{1}{w(Z_{\ell})}\right)$$
$$\leq \frac{1}{n}\mathbb{V}\left(\frac{1}{w(Z_{1})}\right) + 2\frac{1}{n}\sum_{m=1}^{n}\left|\mathbb{C}\left(\frac{1}{w(Z_{0})}, \frac{1}{w(Z_{m})}\right)\right|.$$
(6.24)

Using (2.3) (which implies $\sup_{x \in [0,1]} (1/w(x)) \le C$) and applying the Davydov inequality, we obtain

$$\mathbb{V}\left(\frac{1}{\widehat{\mu}}\right) \le C\frac{1}{n} \left(1 + \sum_{m=1}^{n} a_m^q\right) \le C\frac{1}{n}.$$
(6.25)

The proof of Proposition 6.2 is complete.

Proposition 6.3. For any integer $j \ge \tau$ such that $2^j \le n$ and any $k \in \{0, \ldots, 2^j - 1\}$, let $\alpha_{j,k} = \int_0^1 f(x)\phi_{j,k}(x)dx$ and $\hat{\alpha}_{j,k}$ be (4.2). Then,

(1) there exists a constant C > 0 such that

$$\mathbb{E}\Big(\left(\widehat{\alpha}_{j,k} - \alpha_{j,k}\right)^2\Big) \le C\frac{1}{n}; \tag{6.26}$$

(2) there exists a constant C > 0 such that

$$\mathbb{E}\Big(\left(\widehat{\alpha}_{j,k} - \alpha_{j,k}\right)^4\Big) \le C2^j \frac{1}{n}.$$
(6.27)

These inequalities hold for $\hat{\beta}_{j,k}$ defined by (4.3) instead of $\hat{\alpha}_{j,k}$, and $\beta_{j,k} = \int_0^1 f(x)\psi_{j,k}(x)dx$ instead of $\alpha_{j,k}$.

Proof of Proposition 6.3. (1) Applying Lemma 6.1 and Proposition 6.2, we have

$$\mathbb{E}\left(\left(\widehat{\alpha}_{j,k} - \alpha_{j,k}\right)^{2}\right) \leq C\left(\mathbb{E}\left(\left(\frac{\mu}{n}\sum_{i=1}^{n}\frac{\phi_{j,k}(Z_{i})}{w(Z_{i})} - \alpha_{j,k}\right)^{2}\right) + \mathbb{E}\left(\left(\frac{1}{\widehat{\mu}} - \frac{1}{\mu}\right)^{2}\right)\right)$$

$$= C\left(\mathbb{V}\left(\frac{\mu}{n}\sum_{i=1}^{n}\frac{\phi_{j,k}(Z_{i})}{w(Z_{i})}\right) + \mathbb{V}\left(\frac{1}{\widehat{\mu}}\right)\right) \leq C\frac{1}{n}.$$
(6.28)

(2) We have

$$\left|\widehat{\alpha}_{j,k} - \alpha_{j,k}\right| \le \left|\widehat{\alpha}_{j,k}\right| + \left|\alpha_{j,k}\right|. \tag{6.29}$$

By (2.3), we have $|\hat{\mu}| \le C$ and $\sup_{x \in [0,1]} (1/w(x)) \le C$. So,

$$\left| \frac{\widehat{\mu}}{n} \sum_{i=1}^{n} \frac{\phi_{j,k}(Z_i)}{w(Z_i)} \right| \leq C \frac{1}{n} \sum_{i=1}^{n} \left| \frac{\phi_{j,k}(Z_i)}{w(Z_i)} \right| \leq C \sup_{x \in [0,1]} \left| \frac{\phi_{j,k}(x)}{w(x)} \right|$$

$$\leq C \sup_{x \in [0,1]} \left| \phi_{j,k}(x) \right| \leq C 2^{j/2}.$$
(6.30)

By (6.4), we have $|\alpha_{j,k}| \leq C$. Therefore

$$\left|\hat{\alpha}_{j,k} - \alpha_{j,k}\right| \le C\left(2^{j/2} + 1\right) \le C2^{j/2}.$$
 (6.31)

It follows from (6.31) and (6.28) that

$$\mathbb{E}\Big(\left(\widehat{\alpha}_{j,k} - \alpha_{j,k}\right)^4\Big) \le C2^j \mathbb{E}\Big(\left(\widehat{\alpha}_{j,k} - \alpha_{j,k}\right)^2\Big) \le C2^j \frac{1}{n}.$$
(6.32)

The proof of Proposition 6.3 is complete.

Proposition 6.4. For any $j \in {\tau, ..., j_1}$ and any $k \in {0, ..., 2^j - 1}$, let $\beta_{j,k} = \int_0^1 f(x) \psi_{j,k}(x) dx$, $\hat{\beta}_{j,k}$ be (4.3) and λ_n be (4.8). Then there exist two constants, $\kappa > 0$ and C > 0, such that

$$\mathbb{P}\left(\left|\widehat{\beta}_{j,k} - \beta_{j,k}\right| \ge \frac{\kappa\lambda_n}{2}\right) \le C\frac{1}{n^4}.$$
(6.33)

Proof of Proposition 6.4. It follows from Lemma 6.1 that

$$\mathbb{P}\left(\left|\widehat{\beta}_{j,k} - \beta_{j,k}\right| \ge \frac{\kappa\lambda_n}{2}\right) \le P_1 + P_2,\tag{6.34}$$

where

$$P_{1} = \mathbb{P}\left(\left|\frac{\mu}{n}\sum_{i=1}^{n}\frac{\psi_{j,k}(Z_{i})}{w(Z_{i})} - \beta_{j,k}\right| \ge \kappa C\lambda_{n}\right),$$

$$P_{2} = \mathbb{P}\left(\left|\frac{1}{\hat{\mu}} - \frac{1}{\mu}\right| \ge \kappa C\lambda_{n}\right).$$
(6.35)

In order to bound P_1 and P_2 , let us present a Bernstein inequality for exponentially strongly mixing process. We refer to [25, 26].

Lemma 6.5 (see [25, 26]). Let $\gamma > 0$, c > 0, $\theta > 1$ and $(Z_i)_{i \in \mathbb{Z}}$ be a stationary process such that, for any $m \in \mathbb{Z}$, the associated mth strongly mixing coefficient (2.2) satisfies $a_m \leq \gamma \exp(-c|m|^{\theta})$. Let $n \in \mathbb{N}^*$, $h : \mathbb{R} \to \mathbb{R}$ be a measurable function and, for any $i \in \mathbb{Z}$, $U_i = h(Z_i)$. One assumes that $\mathbb{E}(U_1) = 0$ and there exists a constant M > 0 satisfying $|U_1| \leq M < \infty$. Then, for any $m \in \{1, ..., n\}$ and any $\lambda > 4mM/n$, one has

$$\mathbb{P}\left(\left|\frac{1}{n}\sum_{i=1}^{n}U_{i}\right| \geq \lambda\right) \leq 4\exp\left(-\frac{\lambda^{2}n}{m(64\mathbb{E}(U_{1}^{2})+8\lambda M/3)}\right) + 4\gamma\frac{n}{m}\exp\left(-cm^{\theta}\right).$$
(6.36)

Upper Bound for P_1

For any $i \in \{1, \ldots, n\}$, set

$$U_{i} = \mu \frac{\psi_{j,k}(Z_{i})}{w(Z_{i})} - \beta_{j,k}.$$
(6.37)

Then U_1, \ldots, U_n are identically distributed, depend on the stationary strongly mixing process $(Z_i)_{i \in \mathbb{Z}}$ which satisfies (2.2), Proposition 6.2 gives

$$\mathbb{E}(U_1) = 0, \qquad \mathbb{E}\left(U_1^2\right) \le \mathbb{E}\left(\left(\mu \frac{\psi_{j,k}(Z_1)}{w(Z_1)}\right)^2\right) \le C \tag{6.38}$$

and, by (2.3) and (6.4),

$$|U_{1}| \leq \mu \sup_{x \in [0,1]} \left| \frac{\psi_{j,k}(x)}{w(x)} \right| + \left| \beta_{j,k} \right| \leq C \left(\sup_{x \in [0,1]} \left| \psi_{j,k}(x) \right| + 1 \right)$$

$$\leq C \left(2^{j/2} + 1 \right) \leq C 2^{j/2}.$$
(6.39)

It follows from Lemma 6.5 applied with U_1, \ldots, U_n , $\lambda = \kappa C \lambda_n$, $\lambda_n = ((\ln n)^{1+1/\theta}/n)^{1/2}$, $m = (u \ln n)^{1/\theta}$ with u > 0 (chosen later), $M = C2^{j/2}$ and $2^j \le 2^{j_1} \le n/(\ln n)^{1+1/\theta}$, that

$$P_{1} = \mathbb{P}\left(\left|\frac{1}{n}\sum_{i=1}^{n}U_{i}\right| \geq \kappa C\lambda_{n}\right)$$

$$\leq 4 \exp\left(-C\frac{\kappa^{2}\lambda_{n}^{2}n}{m(1+\kappa\lambda_{n}M)}\right) + 4\gamma\frac{n}{m}\exp\left(-cm^{\theta}\right)$$

$$\leq 4 \exp\left(-C\frac{\kappa^{2}(\ln n)^{1+1/\theta}}{(u\ln n)^{1/\theta}\left(1+\kappa 2^{j/2}\left((\ln n)^{1+1/\theta}/n\right)^{1/2}\right)}\right)$$

$$+ 4\gamma\frac{n}{(u\ln n)^{1/\theta}}\exp\left(-cu\ln n\right)$$

$$\leq C\left(n^{-C\kappa^{2}/(u^{1/\theta}(1+\kappa))} + n^{1-cu}\right).$$
(6.40)

Therefore, for large enough κ and u, we have

$$P_1 \le C \frac{1}{n^4}.\tag{6.41}$$

Upper Bound for P₂

For any $i \in \{1, \ldots, n\}$, set

$$U_i = \frac{1}{w(Z_i)} - \frac{1}{\mu}.$$
 (6.42)

Then U_1, \ldots, U_n are identically distributed, depend on the stationary strongly mixing process $(Z_i)_{i \in \mathbb{Z}}$ which satisfies (2.2), Proposition 6.2 gives

$$\mathbb{E}(U_1) = 0, \qquad \mathbb{E}\left(U_1^2\right) \le \mathbb{E}\left(\frac{1}{\left(w(Z_1)\right)^2}\right) \le C.$$
(6.43)

By (2.3), we have

$$|U_1| \le \sup_{x \in [0,1]} \frac{1}{w(x)} + \left|\frac{1}{\mu}\right| \le C.$$
(6.44)

It follows from Lemma 6.5 applied with U_1, \ldots, U_n , $\lambda = \kappa C \lambda_n$, $\lambda_n = ((\ln n)^{1+1/\theta}/n)^{1/2}$, $m = (u \ln n)^{1/\theta}$ with u > 0 (chosen later) and M = C that

$$P_{2} = \mathbb{P}\left(\left|\frac{1}{n}\sum_{i=1}^{n}U_{i}\right| \geq \kappa C\lambda_{n}\right)$$

$$\leq 4 \exp\left(-C\frac{\kappa^{2}\lambda_{n}^{2}n}{m(1+\kappa\lambda_{n}M)}\right) + 4\gamma\frac{n}{m}\exp\left(-cm^{\theta}\right)$$

$$\leq 4 \exp\left(-C\frac{\kappa^{2}(\ln n)^{1+1/\theta}}{(u\ln n)^{1/\theta}\left(1+\kappa\left((\ln n)^{1+1/\theta}/n\right)^{1/2}\right)}\right)$$

$$+ 4\gamma\frac{n}{(u\ln n)^{1/\theta}}\exp\left(-cu\ln n\right)$$

$$\leq C\left(n^{-C\kappa^{2}/u^{1/\theta}} + n^{1-cu}\right).$$
(6.45)

Therefore, for large enough κ and u, we have

$$P_2 \le C \frac{1}{n^4}.\tag{6.46}$$

Putting (6.34), (6.41), and (6.46) together, this ends the proof of Proposition 6.4. \Box

6.2. Proofs of the Main Results

Proof of Theorem 5.1. We expand the function f on \mathcal{B} as

$$f(x) = \sum_{k=0}^{2^{j_0}-1} \alpha_{j_0,k} \phi_{j_0,k}(x) + \sum_{j=j_0}^{\infty} \sum_{k=0}^{2^{j}-1} \beta_{j,k} \psi_{j,k}(x), \quad x \in [0,1],$$
(6.47)

where $\alpha_{j_0,k} = \int_0^1 f(x)\phi_{j_0,k}(x)dx$ and $\beta_{j,k} = \int_0^1 f(x)\psi_{j,k}(x)dx$. We have, for any $x \in [0, 1]$,

$$\widehat{f}^{L}(x) - f(x) = \sum_{k=0}^{2^{j_0}-1} (\widehat{\alpha}_{j_0,k} - \alpha_{j_0,k}) \phi_{j_0,k}(x) - \sum_{j=j_0}^{\infty} \sum_{k=0}^{2^{j}-1} \beta_{j,k} \psi_{j,k}(x).$$
(6.48)

Since \mathcal{B} is an orthonormal basis of $\mathbb{L}^2([0,1])$, we have,

$$\mathbb{E}\left(\int_{0}^{1} \left(\widehat{f}^{L}(x) - f(x)\right)^{2} dx\right) = \sum_{k=0}^{2^{j_{0}-1}} \mathbb{E}\left(\left(\widehat{\alpha}_{j_{0},k} - \alpha_{j_{0},k}\right)^{2}\right) + \sum_{j=j_{0}}^{\infty} \sum_{k=0}^{2^{j}-1} \beta_{j,k}^{2}.$$
(6.49)

Using Proposition 6.3, we obtain

$$\sum_{k=0}^{2^{j_0}-1} \mathbb{E}\Big(\left(\widehat{\alpha}_{j_0,k} - \alpha_{j_0,k}\right)^2\Big) \le C2^{j_0} \frac{1}{n} \le Cn^{-2s/(2s+1)}.$$
(6.50)

Since $p \ge 2$, we have $B_{p,r}^s(M) \subseteq B_{2,\infty}^s(M)$. Hence

$$\sum_{j=j_0}^{\infty} \sum_{k=0}^{2^j - 1} \beta_{j,k}^2 \le C 2^{-2j_0 s} \le C n^{-2s/(2s+1)}.$$
(6.51)

Therefore,

$$\mathbb{E}\left(\int_0^1 \left(\widehat{f}^L(x) - f(x)\right)^2 dx\right) \le C n^{-2s/(2s+1)}.$$
(6.52)

The proof of Theorem 5.1 is complete.

Proof of Theorem 5.2. We expand the function f on \mathcal{B} as

$$f(x) = \sum_{k=0}^{2^{\tau}-1} \alpha_{\tau,k} \phi_{\tau,k}(x) + \sum_{j=\tau}^{\infty} \sum_{k=0}^{2^{j}-1} \beta_{j,k} \psi_{j,k}(x), \quad x \in [0,1],$$
(6.53)

where $\alpha_{\tau,k} = \int_0^1 f(x)\phi_{\tau,k}(x)dx$ and $\beta_{j,k} = \int_0^1 f(x)\psi_{j,k}(x)dx$. We have, for any $x \in [0, 1]$,

$$\hat{f}^{H}(x) - f(x) = \sum_{k=0}^{2^{\tau}-1} (\hat{\alpha}_{\tau,k} - \alpha_{\tau,k}) \phi_{\tau,k}(x) + \sum_{j=\tau}^{j_{1}} \sum_{k=0}^{2^{j}-1} (\hat{\beta}_{j,k} \mathbb{I}_{\{|\hat{\beta}_{j,k}| \ge \kappa \lambda_{n}\}} - \beta_{j,k}) \psi_{j,k}(x) - \sum_{j=j_{1}+1}^{\infty} \sum_{k=0}^{2^{j}-1} \beta_{j,k} \psi_{j,k}(x).$$
(6.54)

Since \mathcal{B} is an orthonormal basis of $\mathbb{L}^2([0,1])$, we have

$$\mathbb{E}\left(\int_0^1 \left(\widehat{f}^H(x) - f(x)\right)^2 dx\right) = R + S + T,\tag{6.55}$$

where

$$R = \sum_{k=0}^{2^{\tau}-1} \mathbb{E}\Big(\left(\widehat{\alpha}_{\tau,k} - \alpha_{\tau,k}\right)^2\Big), \qquad S = \sum_{j=\tau}^{j_1} \sum_{k=0}^{2^{j}-1} \mathbb{E}\Big(\left(\widehat{\beta}_{j,k} \mathbb{I}_{\{|\widehat{\beta}_{j,k}| \ge \kappa \lambda_n\}} - \beta_{j,k}\right)^2\Big),$$

$$T = \sum_{j=j_1+1}^{\infty} \sum_{k=0}^{2^{j}-1} \beta_{j,k}^2.$$
(6.56)

Let us bound *R*, *T*, and *S*, in turn.

Upper Bound for R

Using Proposition 6.3 and 2s/(2s+1) < 1, we obtain

$$R \le C2^{\tau} \frac{1}{n} \le C \left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2s/(2s+1)}.$$
(6.57)

Upper Bound for T

For $r \ge 1$ and $p \ge 2$, we have $B_{p,r}^s(M) \subseteq B_{2,\infty}^s(M)$. Since 2s/(2s+1) < 2s, we have

$$T \le C \sum_{j=j_1+1}^{\infty} 2^{-2j_s} \le C 2^{-2j_{1s}} \le C \left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2s} \le C \left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2s/(2s+1)}.$$
(6.58)

For *r* ≥ 1 and *p* ∈ [1, 2), we have $B_{p,r}^{s}(M) \subseteq B_{2,\infty}^{s+1/2-1/p}(M)$. Since *s* > 1/*p*, we have *s* + 1/2 − 1/*p* > *s*/(2*s* + 1). So

$$T \le C \sum_{j=j_1+1}^{\infty} 2^{-2j(s+1/2-1/p)} \le C 2^{-2j_1(s+1/2-1/p)}$$

$$\le C \left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2(s+1/2-1/p)} \le C \left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2s/(2s+1)}.$$
(6.59)

Hence, for $r \ge 1$, $\{p \ge 2 \text{ and } s > 0\}$ or $\{p \in [1, 2) \text{ and } s > 1/p\}$, we have

$$T \le C \left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2s/(2s+1)}$$
. (6.60)

Upper Bound for S

Note that we can write the term S as

$$S = S_1 + S_2 + S_3 + S_4, \tag{6.61}$$

where

$$S_{1} = \sum_{j=\tau}^{j_{1}} \sum_{k=0}^{2^{j}-1} \mathbb{E} \left(\left(\widehat{\beta}_{j,k} - \beta_{j,k} \right)^{2} \mathbb{I}_{\{ |\widehat{\beta}_{j,k}| \ge \kappa \lambda_{n} \}} \mathbb{I}_{\{ |\beta_{j,k}| < \kappa \lambda_{n}/2 \}} \right),$$

$$S_{2} = \sum_{j=\tau}^{j_{1}} \sum_{k=0}^{2^{j}-1} \mathbb{E} \left(\left(\widehat{\beta}_{j,k} - \beta_{j,k} \right)^{2} \mathbb{I}_{\{ |\widehat{\beta}_{j,k}| \ge \kappa \lambda_{n} \}} \mathbb{I}_{\{ |\beta_{j,k}| \ge \kappa \lambda_{n}/2 \}} \right),$$

$$S_{3} = \sum_{j=\tau}^{j_{1}} \sum_{k=0}^{2^{j}-1} \mathbb{E} \left(\beta_{j,k}^{2} \mathbb{I}_{\{ |\widehat{\beta}_{j,k}| < \kappa \lambda_{n} \}} \mathbb{I}_{\{ |\beta_{j,k}| \ge 2\kappa \lambda_{n} \}} \right),$$

$$S_{4} = \sum_{j=\tau}^{j_{1}} \sum_{k=0}^{2^{j}-1} \mathbb{E} \left(\beta_{j,k}^{2} \mathbb{I}_{\{ |\widehat{\beta}_{j,k}| < \kappa \lambda_{n} \}} \mathbb{I}_{\{ |\beta_{j,k}| < 2\kappa \lambda_{n} \}} \right).$$
(6.62)

Let us investigate the bounds of S_1 , S_2 , S_3 , and S_4 in turn.

Upper Bounds for S_1 *and* S_3

We have

$$\left\{ \left| \widehat{\beta}_{j,k} \right| < \kappa \lambda_{n}, \left| \beta_{j,k} \right| \ge 2\kappa \lambda_{n} \right\} \subseteq \left\{ \left| \widehat{\beta}_{j,k} - \beta_{j,k} \right| > \frac{\kappa \lambda_{n}}{2} \right\},$$

$$\left\{ \left| \widehat{\beta}_{j,k} \right| \ge \kappa \lambda_{n}, \left| \beta_{j,k} \right| < \frac{\kappa \lambda_{n}}{2} \right\} \subseteq \left\{ \left| \widehat{\beta}_{j,k} - \beta_{j,k} \right| > \frac{\kappa \lambda_{n}}{2} \right\},$$

$$\left\{ \left| \widehat{\beta}_{j,k} \right| < \kappa \lambda_{n}, \left| \beta_{j,k} \right| \ge 2\kappa \lambda_{n} \right\} \subseteq \left\{ \left| \beta_{j,k} \right| \le 2 \left| \widehat{\beta}_{j,k} - \beta_{j,k} \right| \right\}.$$
(6.63)

So,

$$\max(S_1, S_3) \le C \sum_{j=\tau}^{j_1} \sum_{k=0}^{2^j - 1} \mathbb{E}\bigg(\left(\widehat{\beta}_{j,k} - \beta_{j,k} \right)^2 \mathbb{I}_{\{|\widehat{\beta}_{j,k} - \beta_{j,k}| > \kappa \lambda_n/2\}} \bigg).$$
(6.64)

It follows from the Cauchy-Schwarz inequality, Propositions 6.3 and 6.4, and $2^j \le 2^{j_1} \le n$ that

$$\mathbb{E}\left(\left(\widehat{\beta}_{j,k}-\beta_{j,k}\right)^{2}\mathbb{I}_{\left\{|\widehat{\beta}_{j,k}-\beta_{j,k}|>\kappa\lambda_{n}/2\right\}}\right) \leq \left(\mathbb{E}\left(\left(\widehat{\beta}_{j,k}-\beta_{j,k}\right)^{4}\right)\right)^{1/2} \left(\mathbb{P}\left(\left|\widehat{\beta}_{j,k}-\beta_{j,k}\right|>\frac{\kappa\lambda_{n}}{2}\right)\right)^{1/2}$$
$$\leq C\left(2^{j}\frac{1}{n}\right)^{1/2} \left(\frac{1}{n^{4}}\right)^{1/2} \leq C\frac{1}{n^{2}}.$$
(6.65)

Since 2s/(2s+1) < 1, we have

$$\max(S_1, S_3) \le C \frac{1}{n^2} \sum_{j=\tau}^{j_1} 2^j \le C \frac{1}{n^2} 2^{j_1} \le C \frac{1}{n} \le C \left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2s/(2s+1)}.$$
(6.66)

Upper Bound for S_2

Using again Proposition 6.3, we obtain

$$\mathbb{E}\left(\left(\widehat{\beta}_{j,k} - \beta_{j,k}\right)^2\right) \le C\frac{1}{n} \le C\frac{(\ln n)^{1+1/\theta}}{n}.$$
(6.67)

Hence,

$$S_{2} \leq C \frac{(\ln n)^{1+1/\theta}}{n} \sum_{j=\tau}^{j_{1}} \sum_{k=0}^{2^{j}-1} \mathbb{I}_{\{|\beta_{j,k}| > \kappa \lambda_{n}/2\}}.$$
(6.68)

Let j_2 be the integer defined by

$$\frac{1}{2} \left(\frac{n}{\left(\ln n \right)^{1+1/\theta}} \right)^{1/(2s+1)} < 2^{j_2} \le \left(\frac{n}{\left(\ln n \right)^{1+1/\theta}} \right)^{1/(2s+1)}.$$
(6.69)

We have

$$S_2 \le S_{2,1} + S_{2,2},\tag{6.70}$$

where

$$S_{2,1} = C \frac{(\ln n)^{1+1/\theta}}{n} \sum_{j=\tau}^{j_2} \sum_{k=0}^{2^{j-1}} \mathbb{I}_{\{|\beta_{j,k}| > \kappa \lambda_n/2\}},$$

$$S_{2,2} = C \frac{(\ln n)^{1+1/\theta}}{n} \sum_{j=j_2+1}^{j_1} \sum_{k=0}^{2^{j-1}} \mathbb{I}_{\{|\beta_{j,k}| > \kappa \lambda_n/2\}}.$$
(6.71)

We have

$$S_{2,1} \le C \frac{(\ln n)^{1+1/\theta}}{n} \sum_{j=\tau}^{j_2} 2^j \le C \frac{(\ln n)^{1+1/\theta}}{n} 2^{j_2} \le C \left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2s/(2s+1)}.$$
(6.72)

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For $r \ge 1$ and $p \ge 2$, since $B_{p,r}^s(M) \subseteq B_{2,\infty}^s(M)$,

$$S_{2,2} \leq C \frac{(\ln n)^{1+1/\theta}}{n\lambda_n^2} \sum_{j=j_2+1}^{j_1} \sum_{k=0}^{2^{j}-1} \beta_{j,k}^2 \leq C \sum_{j=j_2+1}^{\infty} \sum_{k=0}^{2^{j}-1} \beta_{j,k}^2 \leq C 2^{-2j_2 s}$$

$$\leq C \left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2s/(2s+1)}.$$
(6.73)

For $r \ge 1$, $p \in [1, 2)$ and s > 1/p, using $\mathbb{I}_{\{|\beta_{j,k}| > \kappa \lambda_n/2\}} \le C |\beta_{j,k}|^p / \lambda_n^p, B_{p,r}^s(M) \subseteq B_{2,\infty}^{s+1/2-1/p}(M)$ and (2s+1)(2-p)/2 + (s+1/2-1/p)p = 2s, we have

$$S_{2,2} \leq C \frac{(\ln n)^{1+1/\theta}}{n\lambda_n^p} \sum_{j=j_2+1}^{j_1} \sum_{k=0}^{2^{j}-1} |\beta_{j,k}|^p \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{(2-p)/2} \sum_{j=j_2+1}^{\infty} 2^{-j(s+1/2-1/p)p}$$

$$\leq C \left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{(2-p)/2} 2^{-j_2(s+1/2-1/p)p} \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2s/(2s+1)}.$$
(6.74)

So, for $r \ge 1$, $\{p \ge 2 \text{ and } s > 0\}$ or $\{p \in [1, 2) \text{ and } s > 1/p\}$, we have

$$S_2 \le C \left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2s/(2s+1)}$$
 (6.75)

Upper Bound for S_4

We have

$$S_4 \le \sum_{j=\tau}^{j_1} \sum_{k=0}^{2^j-1} \beta_{j,k}^2 \mathbb{I}_{\{|\beta_{j,k}| \le 2\kappa\lambda_n\}}.$$
(6.76)

Let j_2 be the integer (6.69). Then

$$S_4 \le S_{4,1} + S_{4,2},\tag{6.77}$$

where

$$S_{4,1} = \sum_{j=\tau}^{j_2} \sum_{k=0}^{2^j - 1} \beta_{j,k}^2 \mathbb{I}_{\{|\beta_{j,k}| < 2\kappa\lambda_n\}}, \qquad S_{4,2} = \sum_{j=j_2+1}^{j_1} \sum_{k=0}^{2^j - 1} \beta_{j,k}^2 \mathbb{I}_{\{|\beta_{j,k}| < 2\kappa\lambda_n\}}.$$
(6.78)

We have

$$S_{4,1} \le C \sum_{j=\tau}^{j_2} 2^j \lambda_n^2 = C \frac{(\ln n)^{1+1/\theta}}{n} \sum_{j=\tau}^{j_2} 2^j \le C \frac{(\ln n)^{1+1/\theta}}{n} 2^{j_2} \le C \left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2s/(2s+1)}.$$
 (6.79)

For $r \ge 1$ and $p \ge 2$, since $B_{p,r}^s(M) \subseteq B_{2,\infty}^s(M)$, we have

$$S_{4,2} \le \sum_{j=j_2+1}^{\infty} \sum_{k=0}^{2^{j}-1} \beta_{j,k}^2 \le C 2^{-2j_2 s} \le C \left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2s/(2s+1)}.$$
(6.80)

For $r \ge 1$, $p \in [1, 2)$ and s > 1/p, using $\beta_{j,k}^2 \mathbb{I}_{\{|\beta_{j,k}| < 2\kappa\lambda_n\}} \le C\lambda_n^{2-p} |\beta_{j,k}|^p$, $B_{p,r}^s(M) \subseteq B_{2,\infty}^{s+1/2-1/p}(M)$ and (2s+1)(2-p)/2 + (s+1/2-1/p)p = 2s, we have

$$S_{4,2} \leq C\lambda_n^{2-p} \sum_{j=j_2+1}^{j_1} \sum_{k=0}^{2^{j}-1} |\beta_{j,k}|^p = C\left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{(2-p)/2} \sum_{j=j_2+1}^{j_1} \sum_{k=0}^{2^{j}-1} |\beta_{j,k}|^p$$

$$\leq C\left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{(2-p)/2} \sum_{j=j_2+1}^{\infty} 2^{-j(s+1/2-1/p)p} \leq C\left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2s/(2s+1)}.$$
(6.81)

So, for $r \ge 1$, $\{p \ge 2 \text{ and } s > 0\}$ or $\{p \in [1, 2) \text{ and } s > 1/p\}$, we have

$$S_4 \le C\left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2s/(2s+1)}$$
. (6.82)

It follows from (6.61), (6.66), (6.75), and (6.82) that

$$S \le C \left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2s/(2s+1)}$$
. (6.83)

Combining (6.55), (6.57), (6.60), and (6.83), we have, for $r \ge 1$, $\{p \ge 2 \text{ and } s > 0\}$ or $\{p \in [1, 2) \text{ and } s > 1/p\}$,

$$\mathbb{E}\left(\int_0^1 \left(\widehat{f}^H(x) - f(x)\right)^2 dx\right) \le C\left(\frac{(\ln n)^{1+1/\theta}}{n}\right)^{2s/(2s+1)}.$$
(6.84)

The proof of Theorem 5.2 is complete.

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