

A NOTE ON CONVERGENCE OF WEIGHTED SUMS OF RANDOM VARIABLES

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ABSTRACT. Under uniform integrability condition, some Weak Laws of large numbers are established for weighted sums of random variables generalizing results of Rohatgi, Pruitt and Khintchine. Some Strong Laws of Large Numbers are proved for weighted sums of pairwise independent random variables generalizing results of Jamison, Orey and Pruitt and Etemadi.

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1. INTRODUCTION.

Let X_n , $n \geq 1$ be a sequence of real random variables defined on a probability space (Ω, β, P) and a_{nk} , $n \geq 1$, $k \geq 1$ a double array of real numbers. Limit theorems have been studied in the literature for the sequence $\sum_{k \geq 1} a_{nk} X_k$, $n \geq 1$ of weighted sums of the sequence X_n , $n \geq 1$ under some conditions on the double array of numbers and on the distribution of the sequence X_n , $n \geq 1$. Jamison, Orey and Pruitt [4] studied almost sure convergence of weighted sums under the assumption that the sequence X_n , $n \geq 1$ is independently identically distributed with $E|X_1| < \infty$. One of the objects of this paper is to extend the result of Jamison, Orey and Pruitt [4] on almost sure convergence to cover the case of pairwise independent identically distributed sequences X_n , $n \geq 1$ with $E|X_1| < \infty$. Recently, Etemadi [3] has shown that Strong Law of Large Numbers is valid for sequences X_n , $n \geq 1$ which are pairwise independent identically distributed with $E|X_1| < \infty$. The main result of Section 3 covers Etemadi's result.

Our second objective in this paper is to study convergence in probability of the sequence of weighted sums described above. Convergence in probability has been studied under the following condition.

- (B) There is a random variable X on Ω such that $E|X|^s < \infty$ for some $s \geq 0$ and
- $$P\{|X_n| \geq x\} \leq P\{|X| \geq x\}$$

for every $x \geq 0$ and $n \geq 1$. See Rohatgi [7]. Wei and Taylor [9, Lemma 3, p.284] have shown that if

$$(A) \sup_{n \geq 1} E|X_n|^r < \infty \text{ for some } r > 0$$

holds, then (B) holds for every $0 \leq s < r$. In this paper, we study convergence in probability for sequences of weighted sums under the condition that

$$(C) X_n, n \geq 1 \text{ is uniformly integrable.}$$

One can show that if (B) holds, then $|X_n|^s, n \geq 1$ is uniformly integrable. See Chung [2, Exercise 7, 4.5]. One can also give examples of sequences $X_n, n \geq 1$ satisfying (C) but not (B) with $s = 1$.

2. CONVERGENCE IN PROBABILITY.

In this section, we present some results on convergence of weighted sums from which Weak Law of Large Numbers is derivable. Theorem 1 generalizes some results in the literature in this area. See the remarks following Theorem 1.

THEOREM 1. Let $X_n, n \geq 1$ be a sequence of pairwise independent random variables such that $X_n, n \geq 1$ is uniformly integrable. Let $a_{nk}, n \geq 1, k \geq 1$ be a double array of real numbers satisfying

$$(i) \sum_{k \geq 1} |a_{nk}| \leq C \text{ for every } n \geq 1 \text{ for some constant } C > 0 \text{ and}$$

$$(ii) \max_{k \geq 1} |a_{nk}|, n \geq 1 \text{ converges to } 0.$$

Then $\sum_{k \geq 1} a_{nk}(X_k - EX_k), n \geq 1$ converges to 0 in the mean.

PROOF. It is clear that the series $\sum_{k \geq 1} a_{nk}(X_k - EX_k)$ converges absolutely a.e.

[P] for every $n \geq 1$ since $\sup_{k \geq 1} E|X_k| < \infty$ and $\sum_{k \geq 1} |a_{nk}|$ is convergent. Let $t > 0$. We show

that $\lim_{n \rightarrow \infty} P\{\sum_{k \geq 1} a_{nk}(X_k - EX_k) > t\} = 0$. Let $\epsilon > 0$. Since $X_k, k \geq 1$ is uniformly integrable, there exists $\delta > 0$ such that

$$\sup_{k \geq 1} \int_A |X_k| dP < \epsilon t / 8C \quad (2.1)$$

whenever $A \in \beta$ and $P(A) < \delta$. Further, by Chebychev's inequality, for any $m > 0$ and $k \geq 1$,

$$P\{|X_k| > m\} \leq (1/m)E|X_k| \leq (1/m) \sup_{k \geq 1} E|X_k|.$$

Consequently, there exists a $\gamma > 0$ such that

$$\sup_{k \geq 1} P\{|X_k| > \gamma\} < \delta. \quad (2.2)$$

Define for every $k \geq 1$,

$$Y_k = X_k \quad \text{if } |X_k| \leq \gamma, \\ = 0 \quad \text{otherwise, and}$$

$$Z_k = X_k - Y_k.$$

Note that $Y_k, k \geq 1$ is a sequence of pairwise independent random variables satisfying $|Y_k - EY_k| \leq 2\gamma$ for every $k \geq 1$. By (2.1) and (2.2), we have for every $k \geq 1$,

$$E|Z_k| = \int_{\{|X_k| > \gamma\}} |X_k| dP < \epsilon t / 8C.$$

Consequently, for every $n \geq 1$,

$$E \left| \sum_{k \geq 1} a_{nk} (Z_k - EZ_k) \right| \leq \sum_{k \geq 1} |a_{nk}| E |Z_k - EZ_k| \leq 2 \sum_{k \geq 1} |a_{nk}| E |Z_k| < \epsilon t / 4.$$

Therefore, by Chebychev's inequality, for every $n \geq 1$,

$$P \left\{ \left| \sum_{k \geq 1} a_{nk} (Z_k - EZ_k) \right| > t/2 \right\} < \epsilon / 2. \tag{2.3}$$

Next, we choose $N \geq 1$ such that for every $n \geq N$, we have

$$\max_{k \geq 1} |a_{nk}| < \epsilon t^2 / 32 a^2 C.$$

We observe that for every $n \geq N$,

$$\begin{aligned} P \left\{ \left| \sum_{k \geq 1} a_{nk} (Y_k - EY_k) \right| > t/2 \right\} &\leq (4/t^2) \text{Var} \left(\sum_{k \geq 1} a_{nk} (Y_k - EY_k) \right) \\ &= (4/t^2) \sum_{k \geq 1} a_{nk}^2 E (Y_k - EY_k)^2 \\ &\leq (4/t^2) (\max_{k \geq 1} |a_{nk}|) \sum_{k \geq 1} |a_{nk}| E (Y_k - EY_k)^2 \\ &< \epsilon / 2 \end{aligned} \tag{2.4}$$

Finally, (2.3) and (2.4) yield

$$\begin{aligned} P \left\{ \left| \sum_{k \geq 1} a_{nk} (X_k - EX_k) \right| > t \right\} &\leq P \left\{ \left| \sum_{k \geq 1} a_{nk} (Y_k - EY_k) \right| > t/2 \right\} \\ &\quad + P \left\{ \left| \sum_{k \geq 1} a_{nk} (Z_k - EZ_k) \right| > t/2 \right\} \\ &< \epsilon \quad \text{for every } n \geq N. \end{aligned}$$

Thus $\sum_{k \geq 1} a_{nk} (X_k - EX_k)$, $n \geq 1$ converges to 0 in probability.

To establish mean convergence, it suffices to show that $\sum_{k \geq 1} a_{nk} (X_k - EX_k)$, $n \geq 1$ is uniformly integrable. See Chung [2, Theorem 4.5.4, p.97]. Since X_k , $k \geq 1$ is uniformly integrable and $\sum_{k \geq 1} |a_{nk}| \leq C$ for every $n \geq 1$, it is obvious that $\sum_{k \geq 1} a_{nk} (X_k - EX_k)$, $n \geq 1$ is uniformly integrable.

REMARKS. (1). Rohatgi [7, Theorem 1, p. 305] showed that $\sum_{k \geq 1} a_{nk} (X_k - EX_k)$, $n \geq 1$

converges to 0 in probability under the following conditions.

- (i) X_n , $n \geq 1$ is independent.
- (ii) (B) holds with $s = 1$.
- (iii) The double array a_{nk} , $n \geq 1$, $k \geq 1$ of real numbers satisfies (i) and (ii) of Theorem 1.

In view of the remarks made in the introduction, Theorem 1 generalizes this result of Rohatgi. (Also, this result of Rohatgi was a generalization of a result of Pruitt [6, Theorem 1, p. 770] who started with the assumption that the sequence X_n , $n \geq 1$ is independently identically distributed with $E|X_1| < \infty$). Moreover, our proof is simpler than the one presented by Rohatgi. The essential difference in the proofs lies in the fact that we truncate each X_n at a fixed point a , where as Rohatgi truncated X_n at a_n with a_n varying with n . To illustrate the power of Theorem 1 over Theorem 1 of Rohatgi, consider the following example. Let X_n , $n \geq 1$ be a sequence of pairwise independent random variable with X_n having the following probability law.

$$P\{X_n = n\} = P\{X_n = -n\} = 1/2n\log(n+1),$$

$$P\{X_n = 0\} = 1 - (1/n\log(n+1)).$$

$X_n, n \geq 1$ is uniformly integrable. But (B) does not hold for the sequence $X_n, n \geq 1$ with $s = 1$. Rohatgi's theorem is not applicable to determine the convergence of $(1/n) \sum_{k=1}^n X_k, n \geq 1$ to 0 in probability. But by Theorem 1, the sequence $(1/n) \sum_{k=1}^n X_k,$

$n \geq 1$ does indeed converge to 0 in probability.

(2). Chung [2, Theorem 5.2.2, p. 109] proved (attributed to Khintchine) the result that $(1/n)(X_1 + X_2 + \dots + X_n), n \geq 1$ converges to EX_1 in probability if $X_n, n \geq 1$ is a sequence of pairwise independent identically distributed random variables with $E|X_1| < \infty$. Theorem 1 generalizes this result, Moreover, the proof presented here is much simpler than the one presented by Chung.

If we impose a stronger condition on the double array, we can establish a Weak Law of Large Numbers for weighted sums without the assumption of independence of the random variables but in the presence of uniform integrability.

THEOREM 2. Let $X_n, n \geq 1$ be a sequence of real random variables defined on a probability space (Ω, β, P) such that $|X_n|^r, n \geq 1$ is uniformly integrable for some $0 < r < 1$. Let $a_{nk}, n \geq 1, k \geq 1$ be a double array of real numbers satisfying

- (i) $\sum_{k \geq 1} |a_{nk}|^r \leq C$ for every $n \geq 1$ for some constant $C > 0$,
- (ii) $\max_{k \geq 1} |a_{nk}|, n \geq 1$ converges to 0.

Then $\sum_{k \geq 1} a_{nk} X_k, n \geq 1$ converges to 0 in r-th mean.

PROOF. This can be proved by a simple modification of the proof of Theorem 1. The series $\sum_{k \geq 1} a_{nk} X_k, n \geq 1$ converges absolutely a.e. [P] for every $n \geq 1$ since

$\sum_{k \geq 1} |a_{nk}|^r |X_k|^r$ converges a.e. [P] and $0 < r < 1$. The inequality (2.1) takes the form

$$\sup_{k \geq 1} \int_A |X_k|^r dP < \epsilon t / 8C,$$

and the inequality (2.2) remains intact as it is. The sequences $Y_k, k \geq 1$ and $Z_k, k \geq 1$ are defined in exactly the same way as it was done in the above proof. The probability $P\{\sum_{k \geq 1} a_{nk} Z_k > t/2\}$ is estimated by $\sum_{k \geq 1} |a_{nk}|^r E|Z_k|^r$. Now comes the point of departure.

In the proof of Theorem 1, we showed that $\sum_{k \geq 1} a_{nk} (Y_k - EY_k), k \geq 1$ converges to zero

in probability. Under the conditions of Theorem 2, we can do better than this. The sequence $\sum_{k \geq 1} a_{nk} Y_k, n \geq 1$ does indeed converge to 0 a.e. [P]. This follows from the following chain of inequalities. For every $n \geq 1$,

$$|\sum_{k \geq 1} a_{nk} Y_k| \leq a \sum_{k \geq 1} |a_{nk}| \leq a(\max_{k \geq 1} |a_{nk}|)^{1-r} \sum_{k \geq 1} |a_{nk}|^r.$$

It now follows that $\sum_{k \geq 1} a_{nk} X_k, n \geq 1$ converges to 0 in probability. To establish con-

vergence in r -th mean, it suffices to show that $\left| \sum_{k \geq 1} a_{nk} X_k \right|^r, n \geq 1$ is uniformly integrable. This is not hard to prove.

REMARKS. Rohatgi [7, Theorem 1, p.305] established a weaker conclusion than the one given above under stronger conditions that the sequence $X_n, n \geq 1$ is independently distributed and that (B) holds for the sequence $X_n, n \geq 1$ with $s = r$. The major improvement achieved by Theorem 2 over Rohatgi's theorem is dropping the assumption of independence.

Now, we enquire about the validity of Theorems 1 and 2 in the context of separable Banach spaces. Theorem 2 is valid for sequences of random variables taking values in separable Banach spaces. For the sake of clarity, we give the statement below.

THEOREM 3. Let $X_n, n \geq 1$ be a sequence of random elements taking values in a separable Banach space B equipped with a norm $\|\cdot\|$ such that $\|X_n\|^r, n \geq 1$ is uniformly integrable for some $0 < r < 1$. Let $a_{nk}, n \geq 1, k \geq 1$ be a double array of real numbers satisfying (i) and (ii) of Theorem 2. Then $\sum_{k \geq 1} a_{nk} X_k, n \geq 1$ converges to 0 in the r -th mean, i.e., $E \left| \sum_{k \geq 1} a_{nk} X_k \right|^r, n \geq 1$ converges to zero.

REMARKS. (1). The proof of Theorem 3 is analogous to the one given for Theorem 2. This result is a generalization of Theorem 2.1 of [1]. The major improvement achieved in the above result is in disposing of the assumption of independence in Theorem 2.1 of [1]. Moreover, the proof suggested above is much simpler than the one presented in [1] for Theorem 2.1.

(2). Theorem 1 is not valid for Banach space-valued random variables under conditions similar to those imposed in Theorem 1. (See the comments following Theorem 1.1 of [1]). For the validity of Theorem 3, almost sure convergence of $\sum_{k \geq 1} a_{nk} Y_k, n \geq 1$ to 0 certainly helped. The proof given for Theorem 1 fails to work in Banach spaces because we are unable to establish convergence of $\sum_{k \geq 1} a_{nk} Y_k, n \geq 1$ to 0 either in probability or a.e. [P]. However, if $X_n, n \geq 1$ is uniformly tight, i.e., given $\epsilon > 0$, there exists a compact subset C of B such that $P\{X_n \in C\} > 1 - \epsilon$ for every $n \geq 1$, then Theorem 1 is valid under the conditions stipulated therein. For further details on this result, see Wang and Bhaskara Rao [10, Theorem 2.4].

3. ON STRONG CONVERGENCE.

Extensions of Kolmogorov's Strong Law of Large Numbers are generally sought so that they become more applicable under circumstances less stringent than those imposed by Kolmogorov's Strong Law of Large Numbers. Jamison, Orey and Pruitt [4, Theorem 3, p. 42] worked with independent identically distributed sequences of random variables but imposed conditions on the weights to establish Strong Law of Large Numbers. Etemadi [3, Theorem 1, p. 119] relaxed the assumption of independence in Kolmogorov's Strong Law of Large Numbers to pairwise independence and arrived at the same conclusion. The following result encompasses both these extensions.

THEOREM 4. Let $X_n, n \geq 1$ be a sequence of pairwise independent identically distributed random variables defined on a probability space (Ω, β, P) satisfying $E|X_1| < \infty$. Let $a_n, n \geq 1$ be a sequence of positive numbers satisfying $\lim_{n \rightarrow \infty} \max_{1 \leq i \leq n} (a_i/A_n) = 0$, where

$A_n = \sum_{i=1}^n a_i, n \geq 1$. Let $N(n)$ be the number of positive integers $k \geq 1$ such that $(A_k/a_k) \leq n, n \geq 0$. If $N(n)/n \leq r$ for every $n \geq 1$ for some constant $r > 0$, then $\sum_{k=1}^n (a_k/A_n)X_k, n \geq 1$ converges to EX_1 a.e.[P].

PROOF. The proof is carried out in the following three major steps. 1^o. Since each of the sequences $X_n^+, n \geq 1$ and $X_n^-, n \geq 1$ satisfies the hypothesis of the theorem, we can assume, without loss of generality, that each $X_n \geq 0$. For each $k \geq 1$, let

$$Y_k = X_k \quad \text{if } X_k < A_k/a_k, \\ = 0 \quad \text{otherwise.}$$

Note that $Y_k, k \geq 1$ is a sequence of pairwise independent random variables. It can be shown that $\sum_{n \geq 1} (a_n^2/A_n^2)EY_n^2$ is convergent. Further, $\sum_{k=1}^n (a_k/A_n)X_k, n \geq 1$ converges to EX_1 a.e.[P] if and only if $\sum_{k=1}^n (a_k/A_n)(Y_k - EY_k), n \geq 1$ converges to 0 a.e. [P]. The details are worked out in Stout [8, p. 221].

2^o. We now show that almost sure convergence takes place along some well chosen special subsequences of $\sum_{k=1}^n (a_k/A_n)(Y_k - EY_k), n \geq 1$. In the final step 3^o, we will show that convergence takes place along the entire sequence almost surely. Let $\alpha > 1$. Define a sequence m_1, m_2, \dots of positive integers by letting $m_1 = 1$ and $m_i = \min \{j \geq 1; A_j \geq A_{m_{i-1}}^\alpha\}, i = 2, 3, \dots$. Obviously, $m_1 < m_2 < m_3 < \dots$. We show that

$$\sum_{k=1}^{m_i} (a_k/A_{m_i})(Y_k - EY_k), i \geq 1 \text{ converges to } 0 \text{ a.e. [P].}$$

It suffices to show that for

any $\epsilon > 0$,

$$B_{\alpha, \epsilon} = \sum_{i \geq 1} P\{|\sum_{k=1}^{m_i} (a_k/A_{m_i})(Y_k - EY_k)| \geq \epsilon\} < \infty.$$

By Chebychev's inequality,

$$B_{\alpha, \epsilon} \leq (1/\epsilon^2) \sum_{i \geq 1} \text{Var}(\sum_{k=1}^{m_i} (a_k/A_{m_i})(Y_k - EY_k)) \\ = (1/\epsilon^2) \sum_{i \geq 1} \sum_{k=1}^{m_i} (a_k^2/A_{m_i}^2) \text{Var}(Y_k) \\ \leq (1/\epsilon^2) \sum_{i \geq 1} \sum_{k=1}^{m_i} (a_k^2/A_{m_i}^2) EY_k^2 \\ = (1/\epsilon^2) \sum_{k \geq 1} a_k^2 EY_k^2 \sum_{j \geq j_k} 1/A_{m_j}^2,$$

where $j_k = \min\{j \geq 1; m_j \geq k\}, j = 1, 2, 3, \dots$.

It is clear that $A_{m_{j_k}} \geq A_k$ and $\sum_{j \geq j_k} 1/A_{m_j}^2 \leq (1/A_{m_{j_k}}^2)(1 + 1/\alpha^2 + 1/\alpha^4 + 1/\alpha^6 + \dots)$

$$= (1/A_{m_{j_k}}^2)(\alpha^2/\alpha^2 - 1) \leq (1/A_k^2)(\alpha^2/\alpha^2 - 1).$$

Consequently,

$$B_{\alpha, \epsilon} \leq (1/\epsilon^2)(\alpha^2/\alpha^2 - 1) \sum_{k \geq 1} (a_k^2/A_n^2) EY_k^2 < \infty, \text{ by } 1^o.$$

3^o. Finally we show that the entire sequence $\sum_{k=1}^n (a_k/A_n)(Y_k - EY_k), n \geq 1$ converges

to 0 a.e.[P]. Since $\lim_{n \rightarrow \infty} \max_{1 \leq k \leq n} (a_k/A_n) = 0, \lim_{n \rightarrow \infty} (a_n/A_{n-1}) = 0$ and $\lim_{n \rightarrow \infty} (A_n/A_{n-1}) = 1$.

Observe also that $A_n/A_{n-1} > 1$ for every $n \geq 2$. We can find an integer $N_\alpha \geq 1$ such that $1 < (A_n/A_{n-1}) < \alpha$ for all $n \geq N_\alpha$. Equivalently, if $n \geq N_\alpha$, $A_{n-1} < A_n < \alpha A_{n-1}$. For each $n \geq 1$, there exists a unique integer $i \geq 2$ such that $m_{i-1} \leq n < m_i$. If $m_{i-1} \geq N_\alpha$, then

$$A_n < A_{m_i} < \alpha A_{m_{i-1}} < \alpha^2 A_{m_{i-1}} < \alpha^2 A_n. \tag{3.1}$$

Consequently, $1/A_n < \alpha^2/A_{m_i}$ and $\sum_{k=1}^n (a_k/A_n)Y_k \leq \alpha^2 \sum_{k=1}^{m_i} (a_k/A_{m_i})Y_k$ for every $n \geq N_\alpha$.

Therefore,

$$\limsup_{n \rightarrow \infty} \sum_{k=1}^n (a_k/A_n)Y_k \leq \lim_{i \rightarrow \infty} \sum_{k=1}^{m_i} (a_k/A_{m_i})Y_k = \alpha^2 EX_1 \text{ a.e. [P].}$$

(One can check that $\lim_{n \rightarrow \infty} EY_n = EX_1$ and, by Toeplitz lemma, $\lim_{n \rightarrow \infty} \sum_{k=1}^n (a_k/A_n)EY_k = EX_1$.)

From (3.1), we also observe that $1/A_n > 1/\alpha^2 A_{m_{i-1}}$ and $\sum_{k=1}^n (a_k/A_n)Y_k \geq (1/\alpha^2) \sum_{k=1}^{m_{i-1}}$

$(a_k/A_{m_{i-1}})Y_k$, if $m_{i-1} \geq N_\alpha$. Consequently,

$$\liminf_{n \rightarrow \infty} \sum_{k=1}^n (a_k/A_n)Y_k \geq (1/\alpha^2) \lim_{i \rightarrow \infty} \sum_{k=1}^{m_{i-1}} (a_k/A_{m_{i-1}})Y_k = (1/\alpha^2)EX_1 \text{ a.e. [P].}$$

Thus we observe that for every $\alpha > 1$,

$$(1/\alpha^2)EX_1 \leq \liminf_{n \rightarrow \infty} \sum_{k=1}^n (a_k/A_n)Y_k \leq \limsup_{n \rightarrow \infty} \sum_{k=1}^n (a_k/A_n)Y_k \leq \alpha^2 EX_1 \text{ a.e. [P]. This proves}$$

the almost sure convergence of the desired sequence.

REMARKS. (1). This theorem extends to separable Banach spaces-valued random variables verbatim. A proof can easily be obtained with appropriate modifications of the proof of Strong Law of Large Numbers given in Padgett and Taylor [5, p.42-44]. Or, one can adopt the argument given in Bozorgnia and Bhaskara Rao [1] in the proof of their Theorem 2.1.

(2). Kolmogorov's inequality plays a crucial role in the standard proof offered in many text books for Kolmogorov's Strong Law of Large Numbers. This inequality, as it stands, cannot be commandeered for pairwise independent sequences of random variables. The idea of establishing convergence along some special subsequences is taken from Etemadi [3] but his technique has been modified extensively in the above proof to suit our needs. Also, Theorem 4 strengthens the conclusion of Theorem 5.2.2 of Chung [2, p. 109] from convergence in probability to convergence almost everywhere [P].

There are sequences $a_n, n \geq 1$ of positive numbers such that $\lim_{n \rightarrow \infty} \max_{1 \leq k \leq n} a_k/A_n = 0$

but $N(n)/n, n \geq 1$ is unbounded. See Jamison, Orey and Pruitt [4, p.43]. In such a case, Theorem 4 becomes inapplicable. However, if we impose a stronger condition that $E|X_1| \log^+ |X_1| < \infty$, one can establish a Strong Law of Large Numbers generalizing Theorem 4 of Jamison, Orey and Pruitt [4, p.43] as follows.

THEOREM 5. Let $X_n, n \geq 1$ be a sequence of pairwise independent identically distributed real random variables defined on a probability space (Ω, β, P) with $E|X_1| \log^+ |X_1| < \infty$. Let $a_n, n \geq 1$ be a sequence of positive numbers satisfying $\lim_{n \rightarrow \infty} A_n = \infty$,

where $A_n = \sum_{i=1}^n a_i, n \geq 1$. Then

$$\sum_{k=1}^n (a_k/A_n) X_k, n \geq 1 \text{ converges to } EX_1 \text{ a.e. [P].}$$

PROOF. Using Lemma 2 of Jamison, Orey and Pruitt [4, p.43], one can prove the above result by a suitable modification of the proof of Theorem 4.

REMARK. Theorem 5 is also valid for separable Banach space-valued random variables. The relevant moment condition is that $E\|X_1\|\log^+\|X_1\| < \infty$.

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